

# CS 6210: Matrix Computations

## Rank deficiency and regularization

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### Bias-variance decomposition

The *bias-variance decomposition* is a standard result in statistics. Suppose we want to predict

$$y = f(x) + \epsilon$$

where  $f(x)$  is some underlying “ground truth” function and  $\epsilon$  is a noise term with mean zero and variance  $\sigma^2$ . Some algorithm takes in a data set  $D$  that encompasses our observations about the function (usually  $D = \{(x_i, y_i)\}_{i=1}^n$ ) and produces a prediction function  $s_D(x)$ . The data set (and the algorithm) may include some randomness, which we assume is independent of the measurement noise  $\epsilon$  at any test point. So we also define the mean and variance for the prediction function:

$$\bar{s}_D(x) = \mathbb{E}_D[s_D(x)], \quad \text{Var}_D[s_D(x)] = \mathbb{E}_D[(s_D(x) - \bar{s}_D(x))^2].$$

The *bias* in the predictor is the difference between its mean and the true function  $f(x)$ :

$$\text{Bias}_D[s_D(x)] = f(x) - \bar{s}_D(x).$$

The bias-variance decomposition says that the squared prediction error at a test point  $s$  can be written

$$\mathbb{E}_{D,\epsilon}[(y - s_D(x))^2] = \text{Var}_D[s_D(x)] + \text{Bias}_D[s_D(x)]^2 + \sigma^2.$$

The argument consists of two steps. First, we expand the quadratic and use independence of  $\epsilon$  from  $D$  to get

$$\begin{aligned} \mathbb{E}_{D,\epsilon}[(y - s_D(x))^2] &= \mathbb{E}_D[(f(x) - s_D(x))^2] + 2\mathbb{E}_{D,\epsilon}[\epsilon(f(x) - s_D(x))] + \mathbb{E}_\epsilon[\epsilon^2] \\ &= \mathbb{E}_D[(f(x) - s_D(x))^2] + 2\mathbb{E}_\epsilon[\epsilon]\mathbb{E}_D[(f(x) - s_D(x))] + \mathbb{E}_\epsilon[\epsilon^2] \\ &= \mathbb{E}_D[(f(x) - s_D(x))^2] + \sigma^2 \end{aligned}$$

since  $\mathbb{E}_\epsilon[\epsilon] = 0$  and  $\mathbb{E}_\epsilon[\epsilon^2] = \sigma^2$ . Now write

$$f(x) - s_D(x) = b(x) - e_D(x)$$

where

$$\begin{aligned} b(x) &:= f(x) - \bar{s}_D(x) \\ e_D(x) &:= s_D(x) - \bar{s}_D(x). \end{aligned}$$

Note that  $b(x)$  the bias and that  $\mathbb{E}_D[e_D(x)] = 0$  and  $\mathbb{E}_D[e_D(x)^2] = \text{Var}_D[s_D(x)]$ , so that

$$\begin{aligned} \mathbb{E}_D[(b(x) - e_D(x))^2] &= b(x)^2 - 2b(x)\mathbb{E}_D(x) + \mathbb{E}_D[e_D(x)^2] \\ &= b(x)^2 + \text{Var}_D[s_D(x)^2] \\ &= \text{Bias}_D[s_D(x)]^2 + \text{Var}_D[s_D(x)^2]. \end{aligned}$$

## Bias-variance tradeoffs in the matrix setting

Now consider linear least squares in the context of the bias-variance tradeoff. For simplicity, rather than a function  $f(x)$ , let us assume that our ground truth is a vector of measurements  $f \in \mathbb{R}^M$ . At test time, we are trying to predict  $y = f + e$ , where  $e$  is a vector of mean zero errors. Our model space will be  $\{Ax : x \in \mathbb{R}^n\}$ , and our fitting algorithm will involve

$$\hat{x} = A_1^\dagger y_1, \quad y_1 = f_1 + \tilde{e}_1$$

where  $A_1 \in \mathbb{R}^{m \times n}$  is a subset of the rows of  $A$  and  $f_1$  is the corresponding subset of rows of  $f$ , and  $\tilde{e}_1$  is a vector of mean zero training-time errors. By linearity of expectation, the mean model is

$$\bar{x} = \mathbb{E}[\hat{x}] = A_1^\dagger f_1.$$

and

$$\hat{x} - \bar{x} = A_1^\dagger \tilde{e}_1$$

Then the bias-variance decomposition gives us

$$\mathbb{E}[\|y - A\hat{x}\|^2] = \|f - A\bar{x}\|^2 + \mathbb{E}[\|AA_1^\dagger \tilde{e}_1\|^2] + \mathbb{E}[\|e\|^2].$$

This is all as in the previous section, just with a concrete choice of algorithms for fitting the model.

Now suppose that  $x_* = A^\dagger f$  and  $r = f - Ax_*$ . The model associated with  $x_*$  is optimal: the squared bias term  $\|r\|^2$  is as small as possible, and if a genie gives us  $x_*$ , there is no variance! How does the model associated with  $\hat{x}$  compare? We can rewrite the squared bias term for the  $\hat{x}$  model as

$$\|f - A\bar{x}\|^2 = \|r - A(\bar{x} - x_*)\|^2 = \|r\|^2 + \|A(\bar{x} - x_*)\|^2.$$

Note that  $x_* = A_1^\dagger(b_1 + r_1)$ , so  $\bar{x} - x_* = -A_1^\dagger r_1$ . Therefore

$$\|f - A\bar{x}\|^2 = \|r\|^2 + \|AA_1^\dagger r_1\|^2.$$

Putting things together, we have the decomposition

$$\mathbb{E}[\|y - A\hat{x}\|^2] = (\|r\|^2 + \|AA_1^\dagger r_1\|^2) + (\mathbb{E}[\|AA_1^\dagger \tilde{e}_1\|^2] + \mathbb{E}[\|e\|^2]),$$

where the first two terms come from the bias and the second two terms come from the variance of the training error  $\tilde{e}_1$  and the test error  $e$ .

Note that if  $e$  and  $\tilde{e}$  have independent entries with mean  $\sigma^2$ , then

$$\begin{aligned}\mathbb{E}[\|AA_1^\dagger \tilde{e}_1\|^2] &= \sigma^2 \|AA_1^\dagger\|_F^2 \leq \sigma^2 n \|AA_1^\dagger\|_2^2 \\ \mathbb{E}[\|e\|^2] &= \sigma^2 M\end{aligned}$$

In this case, we have

$$\mathbb{E}[\|y - A\hat{x}\|^2] \leq \left(1 + \|AA_1^\dagger\|^2 \frac{\|r_1\|^2}{\|r\|^2}\right) \|r\|^2 + \left(1 + \|AA_1^\dagger\|^2 \frac{n}{M}\right) M\sigma^2$$

We note that  $\|r_1\|^2/\|r\|^2$  should be approximately  $n/M$  if the residual entries in  $r_1$  are “typical” of those in  $r$ . We can be cruder in our bounds and say

$$\mathbb{E}[\|y - A\hat{x}\|^2] \leq (1 + \|AA_1^\dagger\|^2) (\|r\|^2 + M\sigma^2).$$

This can be interpreted as a *quasi-optimality* result: the expected squared prediction error is within a constant factor  $(1 + \|AA_1^\dagger\|^2)$  of the best possible error given the model flexibility and the noise.

When  $\|A_1^\dagger\|$  is large, the problem of fitting to training data is ill-posed, and the accuracy can be compromised. What can we do? As we discussed in the last section, the problem with ill-posed problems is that they admit many solutions of very similar quality. In order to distinguish between these possible solutions to find a model with good predictive power, we consider *regularization*: that is, we assume that the coefficient vector  $x$  is not too large in norm, or that it is sparse. Different statistical assumptions give rise to different regularization strategies; for the current discussion, we shall focus on the computational properties of a few of the more common regularization strategies without going into the details of the statistical assumptions. In particular, we consider four strategies in turn

1. *Factor selection via pivoted QR.*
2. *Tikhonov regularization* and its solution.
3. *Truncated SVD regularization.*
4.  $\ell^1$  *regularization* or the *lasso*.

## Factor selection and pivoted QR

In ill-conditioned problems, the columns of  $A$  are nearly linearly dependent; we can effectively predict some columns as linear combinations of other columns. The goal of the column pivoted QR algorithm is to find a set of columns that are “as linearly independent as possible.” This is not such a simple task, and so we settle for a greedy strategy: at each step, we select the column that is least well predicted (in the sense of residual norm) by columns already selected. This leads to the *pivoted QR factorization*

$$A\Pi = QR$$

where  $\Pi$  is a permutation and the diagonal entries of  $R$  appear in descending order (i.e.  $r_{11} \geq r_{22} \geq \dots$ ). To decide on how many factors to keep in the factorization, we either automatically take the first  $k$  or we dynamically choose to take  $k$  factors where  $r_{kk}$  is greater than some tolerance and  $r_{k+1,k+1}$  is not.

The pivoted QR approach has a few advantages. It yields *parsimonious* models that predict from a subset of the columns of  $A$  – that is, we need to measure fewer than  $n$  factors to produce an entry of  $b$  in a new column. It can also be computed relatively cheaply, even for large matrices that may be sparse. However, pivoted QR is not the only approach! A related approach due to Golub, Klema, and Stewart computes  $A = U\Sigma V^T$  and chooses a subset of the factors based on pivoted QR of  $V^T$ . More generally, approaches such as the lasso yield an automatic factor selection.

## Tikhonov regularization (ridge regression)

Another approach is to say that we want a model in which the coefficients are not too large. To accomplish this, we add a penalty term to the usual least squares problem:

$$\text{minimize } \|Ax - b\|^2 + \lambda^2 \|x\|^2.$$

Equivalently, we can write

$$\text{minimize } \left\| \begin{bmatrix} A \\ \lambda I \end{bmatrix} x - \begin{bmatrix} b \\ 0 \end{bmatrix} \right\|^2,$$

which leads to the regularized version of the normal equations

$$(A^T A + \lambda^2 I)x = A^T b.$$

In some cases, we may want to regularize with a more general norm  $\|x\|_M^2 = x^T M x$  where  $M$  is symmetric and positive definite, which leads to the regularized equations

$$(A^T A + \lambda^2 M)x = A^T b.$$

If we want to incorporate prior information that pushes  $x$  toward some initial guess  $x_0$ , we may pose the least squares problem in terms of  $z = x - x_0$  and use some form of Tikhonov regularization. If we know of no particular problem structure in advance, the standard choice of  $M = I$  is a good default.

It is useful to compare the usual least squares solution to the regularized solution via the SVD. If  $A = U\Sigma V^T$  is the economy SVD, then

$$\begin{aligned}x_{LS} &= V\Sigma^{-1}U^Tb \\ x_{Tik} &= Vf(\Sigma)^{-1}U^Tb\end{aligned}$$

where

$$f(\sigma) = \frac{1}{\sqrt{\sigma^{-1} + \lambda^2}}.$$

This *filter* of the inverse singular values affects the larger singular values only slightly, but damps the effect of very small singular values.

## Truncated SVD

The Tikhonov filter reduces the effect of small singular values on the solution, but it does not eliminate that effect. By contrast, the *truncated SVD* approach uses the filter

$$f(z) = \begin{cases} z, & z > \sigma_{\min} \\ \infty, & \text{otherwise.} \end{cases}$$

In other words, in the truncated SVD approach, we use

$$x = V_k \Sigma_k^{-1} U_k^T b$$

where  $U_k$  and  $V_k$  represent the leading  $k$  columns of  $U$  and  $V$ , respectively, while  $\Sigma_k$  is the diagonal matrix consisting of the  $k$  largest singular values.

## $\ell^1$ and the lasso

An alternative to Tikhonov regularization (based on a Euclidean norm of the coefficient vector) is an  $\ell^1$  regularized problem

$$\text{minimize } \|Ax - b\|^2 + \lambda \|x\|_1.$$

This is sometimes known as the “lasso” approach. The  $\ell^1$  regularized problem has the property that the solutions tend to become sparse as  $\lambda$  becomes larger. That is, the  $\ell^1$  regularization effectively imposes a factor selection process like that we saw in the pivoted QR approach.

Unlike the pivoted QR approach, however, the  $\ell^1$  regularized solution cannot be computed by one of the standard factorizations of numerical linear algebra. Instead, one treats it as a more general *convex optimization* problem. We will discuss some approaches to the solution of such problems later in the semester.