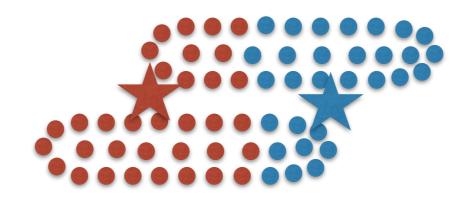
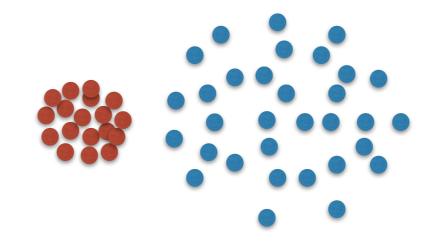
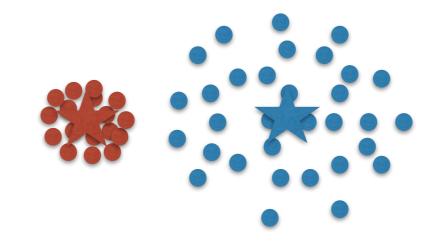
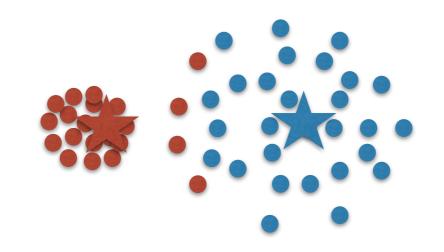
Machine Learning for Data Science (CS4786) Lecture 14

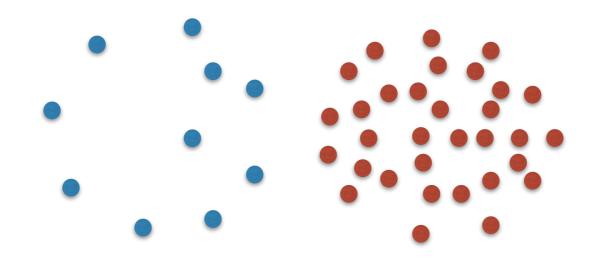
Gaussian Mixture Model

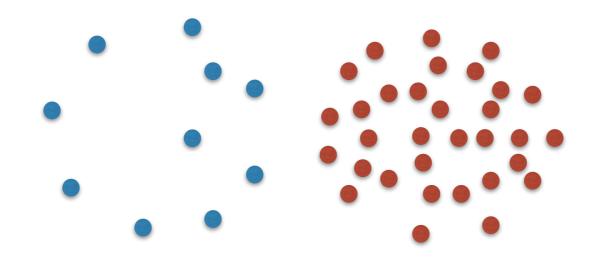


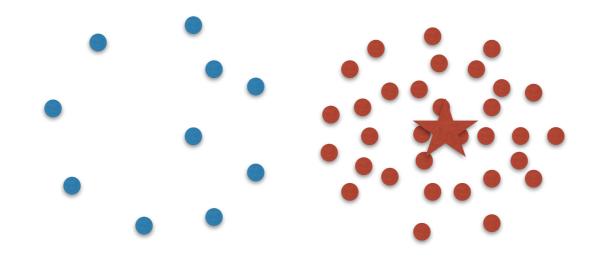


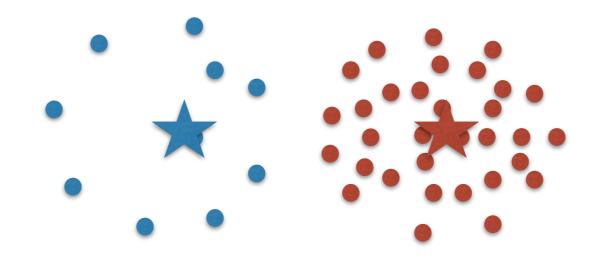


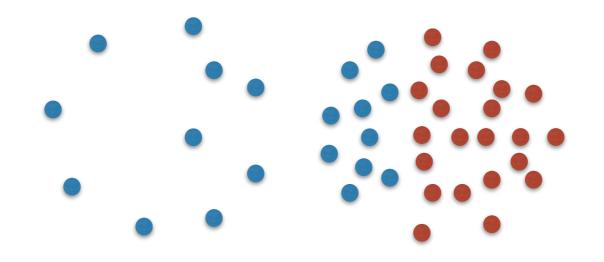


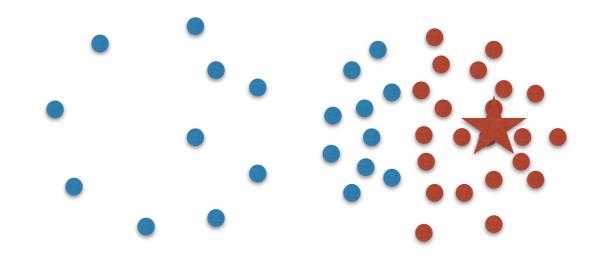


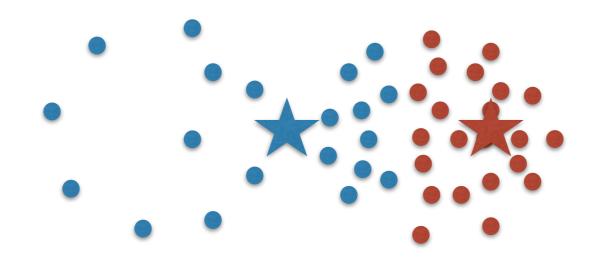












- Of same radius
- Looks for spherical clusters
- And with roughly equal number of points

 Can we design algorithm that can address these shortcomings?





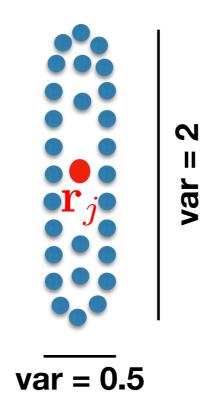


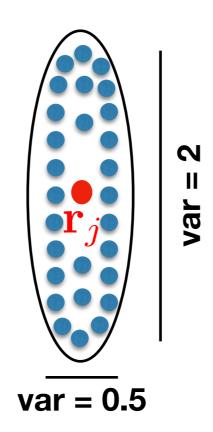
Distance to mean 1 should be smaller than distance to mean 2 as black dot is more likely in cluster 1 than 2

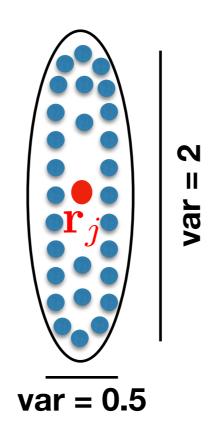


Distance to mean 1 should be smaller than distance to mean 2 as black dot is more likely in cluster 1 than 2

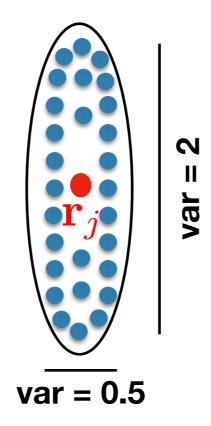
$$d^2(x, C_j) = \frac{(x - \mu_j)^2}{\sigma_j^2}$$



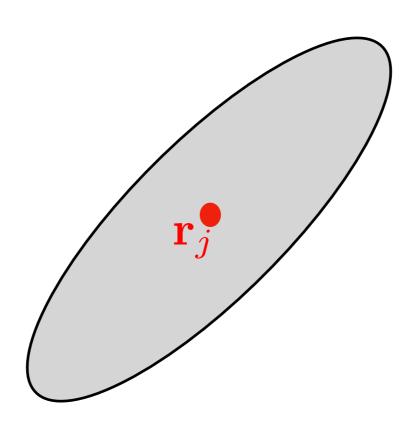




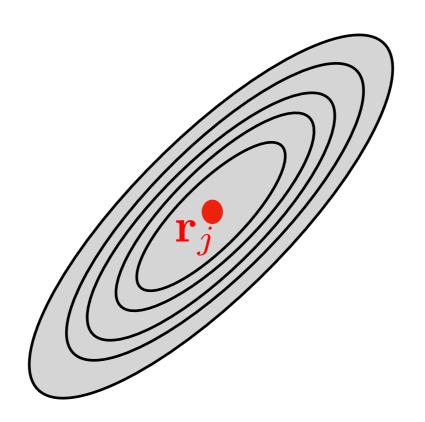
$$(\mathbf{x} - \mathbf{r}_j)^{\top} \begin{bmatrix} 1/0.5 & 0 \\ 0 & 1/2 \end{bmatrix} (\mathbf{x} - \mathbf{r}_j)$$



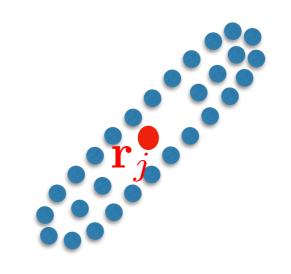
$$(\mathbf{x} - \mathbf{r}_j)^{\top} \left[\sum_{j=1}^{\infty} \mathbf{1} \right] (\mathbf{x} - \mathbf{r}_j)$$

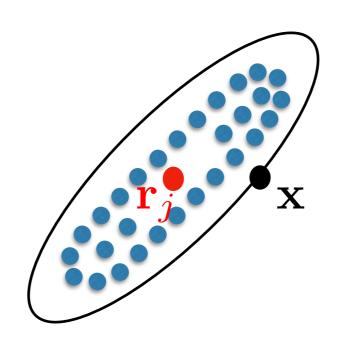


$$(\mathbf{x} - \mathbf{r}_j)^{\top} A(\mathbf{x} - \mathbf{r}_j) \le 1$$

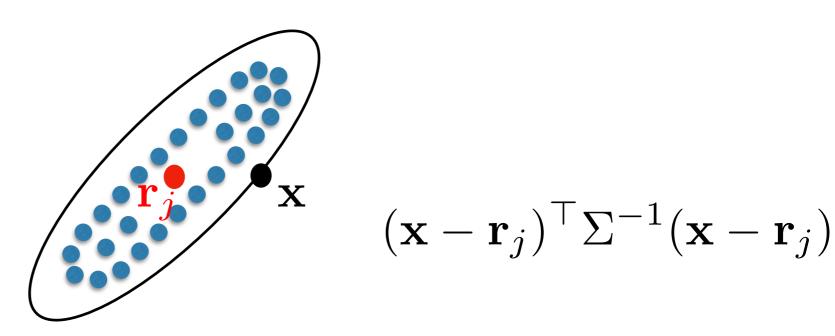


$$(\mathbf{x} - \mathbf{r}_j)^{\top} A(\mathbf{x} - \mathbf{r}_j) \le 1$$





$$(\mathbf{x} - \mathbf{r}_j)^{\top} \Sigma^{-1} (\mathbf{x} - \mathbf{r}_j)$$



$$(\mathbf{x} - \mathbf{r}_j)^{\top} \Sigma^{-1} (\mathbf{x} - \mathbf{r}_j)$$

$$\Sigma = \frac{1}{|C_j|} \sum_{t \in C_j} (\mathbf{x}_t - \mathbf{r}_j) (\mathbf{x}_t - \mathbf{r}_j)^{\top}$$

K-MEANS CLUSTERING

- For all $j \in [K]$, initialize cluster centroids $\hat{\mathbf{r}}_j^0$ randomly and set m = 1
- Repeat until convergence (or until patience runs out)
 - ① For each $t \in \{1, ..., n\}$, set cluster identity of the point

$$\hat{c}^m(\mathbf{x}_t) = \underset{j \in [K]}{\operatorname{argmin}} \|\mathbf{x}_t - \hat{\mathbf{r}}_j^{m-1}\|$$

② For each $j \in [K]$, set new representative as

$$\hat{\mathbf{r}}_j^m = \frac{1}{|\hat{C}_j^m|} \sum_{\mathbf{x}_t \in \hat{C}_i^m} \mathbf{x}_t$$

 $3 m \leftarrow m + 1$

Ellipsoidal Clustering

- For all $j \in [K]$, initialize cluster centroids $\hat{\mathbf{r}}_{j}^{0}$ and ellipsoids $\hat{\Sigma}_{j}^{0}$ randomly and set m = 1
- Repeat until convergence (or until patience runs out)
 - ① For each $t \in \{1, ..., n\}$, set cluster identity of the point

$$\hat{c}^{m}(\mathbf{x}_{t}) = \underset{j \in [K]}{\operatorname{argmin}} \quad (\mathbf{x}_{t} - \hat{\mathbf{r}}_{j}^{m-1})^{\top} \left(\hat{\Sigma}^{m-1}\right)^{-1} \left(\mathbf{x}_{t} - \hat{\mathbf{r}}_{j}^{m-1}\right)$$

② For each $j \in [K]$, set new representative as

$$\hat{\mathbf{r}}_j^m = \frac{1}{|\hat{C}_j^m|} \sum_{\mathbf{x}_t \in \hat{C}_j^m} \mathbf{x}_t \qquad \hat{\Sigma}^m = \frac{1}{|C_j|} \sum_{t \in C_j} (\mathbf{x}_t - \hat{\mathbf{r}}_j^m) (\mathbf{x}_t - \hat{\mathbf{r}}_j^m)^{\top}$$

Ellipsoidal Clustering

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 - ① For each $t \in \{1, ..., n\}$, set cluster identity of the point

$$\hat{c}^{m}(\mathbf{x}_{t}) = \underset{j \in [K]}{\operatorname{argmin}} \quad (\mathbf{x}_{t} - \hat{\mathbf{r}}_{j}^{m-1})^{\mathsf{T}} (\hat{\Sigma}^{m-1})^{-1} (\mathbf{x}_{t} - \hat{\mathbf{r}}_{j}^{m-1})$$

$$d(\mathbf{x}_{t}, C_{j})$$

② For each $j \in [K]$, set new representative as

$$\hat{\mathbf{r}}_j^m = \frac{1}{|\hat{C}_j^m|} \sum_{\mathbf{x}_t \in \hat{C}_j^m} \mathbf{x}_t \qquad \hat{\Sigma}^m = \frac{1}{|C_j|} \sum_{t \in C_j} (\mathbf{x}_t - \hat{\mathbf{r}}_j^m) (\mathbf{x}_t - \hat{\mathbf{r}}_j^m)^{\mathsf{T}}$$

 $m \leftarrow m + 1$

- Looks for spherical clusters
- Of same radius
- And with roughly equal number of points

Looks for spherical clusters



- Of same radius
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- Looks for spherical clusters
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HARD GAUSSIAN MIXTURE MODEL

- For all $j \in [K]$, initialize cluster centroids $\hat{\mathbf{r}}_{j}^{0}$, ellipsoids $\hat{\Sigma}_{j}^{0}$ and initial proportions π^{0} randomly and set m = 1
- Repeat until convergence (or until patience runs out)
 - ① For each $t \in \{1, ..., n\}$, set cluster identity of the point

$$\hat{c}^m(\mathbf{x}_t) = \underset{j \in [K]}{\operatorname{argmin}} \quad (\mathbf{x}_t - \hat{\mathbf{r}}_j^{m-1})^{\top} \left(\hat{\mathbf{\Sigma}}^{m-1}\right)^{-1} (\mathbf{x}_t - \hat{\mathbf{r}}_j^{m-1}) - \log(\pi_j^{m-1})$$

② For each $j \in [K]$, set new representative as

$$\hat{\mathbf{r}}_j^m = \frac{1}{|\hat{C}_j^m|} \sum_{\mathbf{x}_t \in \hat{C}_j^m} \mathbf{x}_t \qquad \hat{\Sigma}^m = \frac{1}{|C_j|} \sum_{t \in C_j} (\mathbf{x}_t - \hat{\mathbf{r}}_j^m) (\mathbf{x}_t - \hat{\mathbf{r}}_j^m)^{\top} \qquad \pi_j^m = \frac{|C_j^m|}{n}$$

 $3 m \leftarrow m + 1$

HARD GAUSSIAN MIXTURE MODEL

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$$\hat{c}^m(\mathbf{x}_t) = \underset{j \in [K]}{\operatorname{argmin}} \quad (\mathbf{x}_t - \hat{\mathbf{r}}_j^{m-1})^{\top} \left(\hat{\mathbf{\Sigma}}^{m-1}\right)^{-1} \left(\mathbf{x}_t - \hat{\mathbf{r}}_j^{m-1}\right) - \log(\pi_j^{m-1})$$

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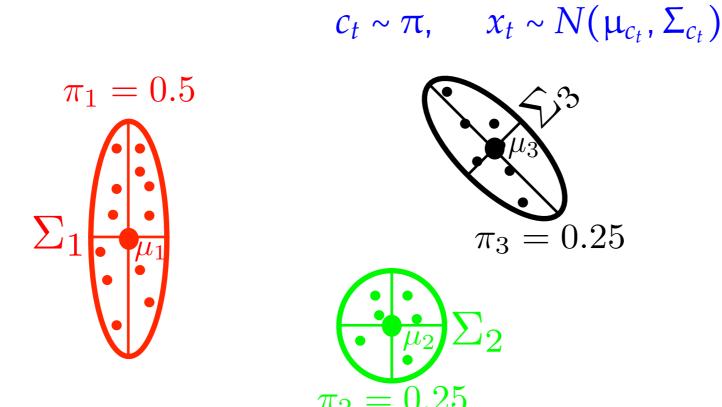
$$\hat{\mathbf{r}}_j^m = \frac{1}{|\hat{C}_j^m|} \sum_{\mathbf{x}_t \in \hat{C}_j^m} \mathbf{x}_t \qquad \hat{\Sigma}^m = \frac{1}{|C_j|} \sum_{t \in C_j} (\mathbf{x}_t - \hat{\mathbf{r}}_j^m) (\mathbf{x}_t - \hat{\mathbf{r}}_j^m)^{\top} \qquad \pi_j^m = \frac{|C_j^m|}{n}$$

 $m \leftarrow m + 1$

Gaussian Mixture Models

Each $\theta \in \Theta$ is a model.

- Gaussian Mixture Model
 - Each θ consists of mixture distribution $\pi = (\pi_1, \dots, \pi_K)$, means $\mu_1, \dots, \mu_K \in \mathbb{R}^d$ and covariance matrices $\Sigma_1, \dots, \Sigma_K$
 - For each t, independently:



Multivariate Gaussian

- Two parameters:
 - Mean $\mu \in \mathbb{R}^d$
 - Covariance matrix \sum of size dxd

$$p(x; \mu, \Sigma) = (2\pi)^{d/2} \det(\Sigma)^{-1/2} \exp\left(-\frac{1}{2}(x - \mu)^{\top} \Sigma^{-1}(x - \mu)\right)$$

Multivariate Gaussian

- Two parameters:
 - Mean $\mu \in \mathbb{R}^d$
 - Covariance matrix Σ of size dxd

$$p(x; \mu, \Sigma) = (2\pi)^{d/2} \det(\Sigma)^{-1/2} \exp\left(-\frac{1}{2}(x - \mu)^{\top} \Sigma^{-1}(x - \mu)\right)$$

0.6

0.4

0.2

PROBABILISTIC MODELS

- consists of set of possible parameters
- We have a distribution P_{θ} over the data induced by each $\theta \in \Theta$
- Data is generated by one of the $\theta \in \Theta$
- Learning: Estimate value or distribution for $\theta^* \in \Theta$ given data

MAXIMUM LIKELIHOOD PRINCIPAL

Pick $\theta \in \Theta$ that maximizes probability of observation

$$\theta_{MLE} = \operatorname{argmax}_{\theta \in \Theta} \log P_{\theta}(x_1, \dots, x_n)$$
Likelihood

EXAMPLE: GAUSSIAN MIXTURE MODEL

MLE:
$$\theta = (\mu_1, \ldots, \mu_K), \pi, \Sigma$$

$$P_{\theta}(x_1, \dots, x_n) = \prod_{t=1}^{n} \left(\sum_{i=1}^{K} \pi_i \frac{1}{\sqrt{(2*3.1415)^2 |\Sigma_i|}} \exp\left(-(x_t - \mu_i)^{\top} \Sigma_i (x_t - \mu_i)\right) \right)$$

Find θ that maximizes $\log P_{\theta}(x_1, \ldots, x_n)$

MLE FOR GMM

Let us consider the one dimensional case, assume variances are 1 and π is uniform

$$\log P_{\theta}(x_{1,...,n}) = \sum_{t=1}^{n} \log \left(\frac{1}{K} \sum_{i=1}^{K} \frac{1}{\sqrt{2 * 3.1415}} \exp\left(-(x_{t} - \mu_{i})^{2} / 2\right) \right)$$

Now consider the partial derivative w.r.t. μ_1 , we have:

$$\frac{\partial \log P_{\theta}(x_{1,...,n})}{\partial \mu_{1}} = \sum_{t=1}^{n} \frac{-(x_{t} - \mu_{1}) \exp\left(-\frac{(x_{t} - \mu_{1})^{2}}{2}\right)}{\sum_{i=1}^{K} \exp\left(-\frac{(x_{t} - \mu_{i})^{2}}{2}\right)}$$

Given all other parameters, optimizing w.r.t. even just μ_1 is hard!

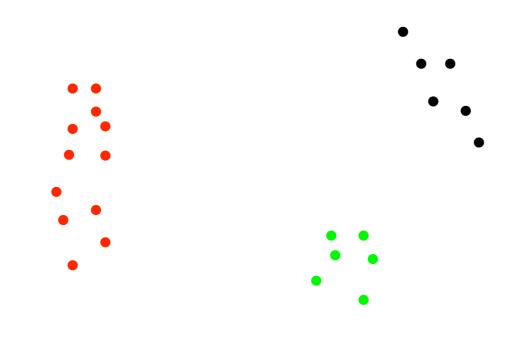
MLE FOR GMM

Say by some magic you knew cluster assignments, then

How would you compute parameters?

MLE FOR GMM

Say by some magic you knew cluster assignments, then



How would you compute parameters?

LATENT VARIABLES

- We only observe x_1, \ldots, x_n , cluster assignments c_1, \ldots, c_n are not observed
- Finding $\theta \in \Theta$ (even for 1-d GMM) that directly maximizes Likelihood or A Posteriori given x_1, \ldots, x_n is hard!
- Given latent variables c_1, \ldots, c_n , the problem of maximizing likelihood (or a posteriori) became easy

Can we use latent variables to device an algorithm?

TOWARDS EM ALGORITHM

• Latent variables can help, but we have a chicken and egg problem

Given all variables including latent variables, finding optimal parameters is easy

Given model parameter, optimizing/finding distribution over the latent variables is easy

HARD GAUSSIAN MIXTURE MODEL

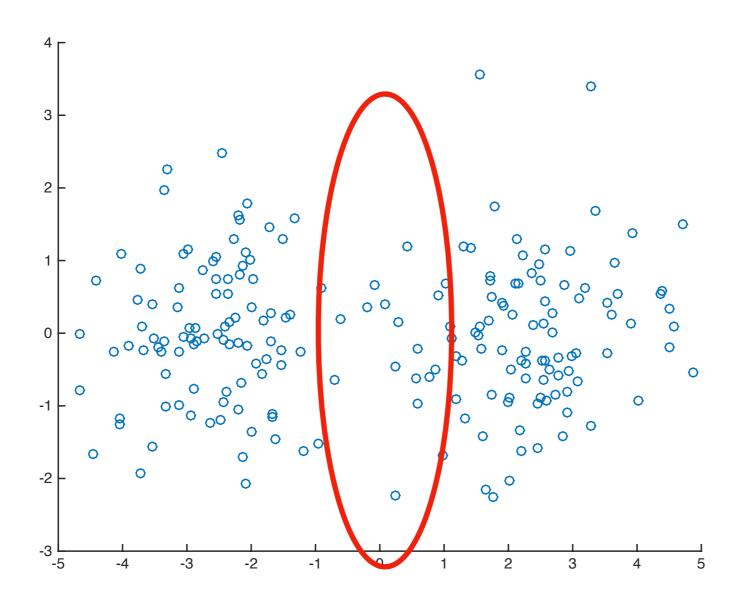
- For all $j \in [K]$, initialize cluster centroids $\hat{\mathbf{r}}_{j}^{0}$, ellipsoids $\hat{\Sigma}_{j}^{0}$ and initial proportions π^{0} randomly and set m = 1
- Repeat until convergence (or until patience runs out)
 - ① For each $t \in \{1, ..., n\}$, set cluster identity of the point

$$\hat{c}^m(\mathbf{x}_t) = \arg\max_{j \in [K]} p(\mathbf{x}_t, \hat{\mathbf{r}}_j^{m-1}, \hat{\Sigma}^{m-1}) \times \pi^m(j)$$

2 For each $j \in [K]$, set new representative as

$$\hat{\mathbf{r}}_j^m = \frac{1}{|\hat{C}_j^m|} \sum_{\mathbf{x}_t \in \hat{C}_j^m} \mathbf{x}_t \qquad \hat{\Sigma}^m = \frac{1}{|C_j|} \sum_{t \in C_j} (\mathbf{x}_t - \hat{\mathbf{r}}_j^m) (\mathbf{x}_t - \hat{\mathbf{r}}_j^m)^\top \qquad \pi_j^m = \frac{|C_j^m|}{n}$$

Pitfall of Hard Assignment



(SOFT) GAUSSIAN MIXTURE MODEL

- For all $j \in [K]$, initialize cluster centroids $\hat{\mathbf{r}}_{j}^{0}$ and ellipsoids $\hat{\Sigma}_{j}^{0}$ randomly and set m = 1
- Repeat until convergence (or until patience runs out)
 - ① For each $t \in \{1, ..., n\}$, set cluster identity of the point

$$Q_t^m(j) = p(\mathbf{x}_t, \hat{\mathbf{r}}_j^{m-1}, \hat{\Sigma}^{m-1}) \times \pi^m(j)$$

② For each $j \in [K]$, set new representative as

$$\hat{\mathbf{r}}_j^m = \frac{\sum_{t=1}^n Q_t(j)\mathbf{x}_t}{\sum_{t=1}^n Q_t(j)} \qquad \hat{\Sigma}^m = \frac{\sum_{t=1}^n Q_t(j)(\mathbf{x}_t - \hat{\mathbf{r}}_j^m)(\mathbf{x}_t - \hat{\mathbf{r}}_j^m)^\top}{\sum_{t=1}^n Q_t(j)}$$

$$\pi_j^m = \frac{\sum_{t=1}^n Q_t(j)}{n}$$

 $m \leftarrow m + 1$