

Lecture 6: Monte Carlo Integration

Chapter 3 in Advanced GI

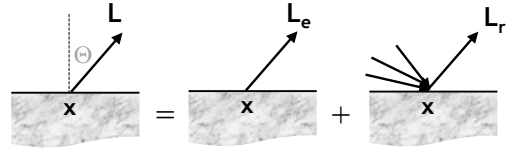
Fall 2004

Kavita Bala

Computer Science

Cornell University

Rendering Equation

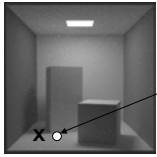


$$L(x \rightarrow \Theta) = L_e(x \rightarrow \Theta) + \int_{\text{hemisphere}} L(x \leftarrow \Psi) f_r(x, \Psi \leftrightarrow \Theta) \cos(\mathbf{N}_x, \Psi) d\omega_\Psi$$

- Applicable for each wavelength

© Kavita Bala, Computer Science, Cornell University

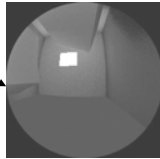
Rendering Equation



$$L(x \rightarrow \Theta) = L_e(x \rightarrow \Theta) +$$

$$\int_{\text{hemisphere}} L(x \leftarrow \Psi) f_r(x, \Psi \leftrightarrow \Theta) \cos(\mathbf{N}_x, \Psi) d\omega_\Psi$$

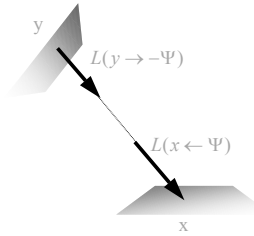
incoming radiance



© Kavita Bala, Computer Science, Cornell University

RE: Area Formulation

$$L(x \rightarrow \Theta) = L_e(x \rightarrow \Theta) + \int_{\Omega_x} f_r(\Psi \leftrightarrow \Theta) \cdot L(x \leftarrow \Psi) \cdot \cos \theta_x \cdot d\omega_\Psi$$



Ray-casting function: what is the nearest visible surface point seen from x in direction Psi?

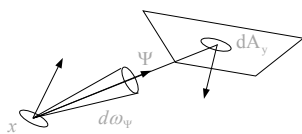
$$y = vp(x, \Psi)$$

$$L(x \leftarrow \Psi) = L(vp(x, \Psi) \rightarrow -\Psi)$$

© Kavita Bala, Computer Science, Cornell University

Rendering Equation

$$L(x \rightarrow \Theta) = L_e(x \rightarrow \Theta) + \int_{\Omega_x} f_r(\Psi \leftrightarrow \Theta) \cdot L(x \leftarrow \Psi) \cdot \cos \theta_x \cdot d\omega_\Psi$$



$$y = vp(x, \Psi)$$

$$L(x \leftarrow \Psi) = L(vp(x, \Psi) \rightarrow -\Psi)$$

$$d\omega_\Psi = \frac{dA_y \cos \theta_y}{r_{xy}^2}$$

© Kavita Bala, Computer Science, Cornell University

Rendering Equation: visible surfaces

$$L(x \rightarrow \Theta) = L_e(x \rightarrow \Theta) + \int_{\Omega_x} f_r(\Psi \leftrightarrow \Theta) \cdot L(x \leftarrow \Psi) \cdot \cos \theta_x \cdot d\omega_\Psi$$

Coordinate transform

$$L(x \rightarrow \Theta) = L_e(x \rightarrow \Theta) + \int_{\text{all surfaces}} f_r(\Psi \leftrightarrow \Theta) \cdot L(y \rightarrow -\Psi) \cos \theta_x \cdot \frac{\cos \theta_y}{r_{xy}^2} \cdot dA_y$$

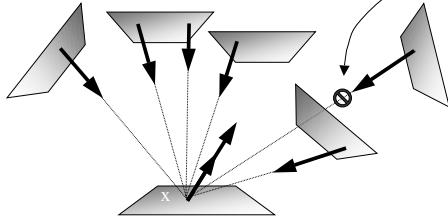
Integration domain = visible surface points y

- Integration domain extended to ALL surface points by including visibility function

© Kavita Bala, Computer Science, Cornell University

Rendering Equation: all surfaces

$$L(x \rightarrow \Theta) = L_e(\dots) + \int_A f_r(\dots) \cdot L(y \rightarrow -\Psi) \cdot \frac{\cos \theta_x \cdot \cos \theta_y}{r_{xy}^2} V(x, y) dA_y$$



© Kavita Bala, Computer Science, Cornell University

Two forms of the RE

- Hemisphere integration

$$L(x \rightarrow \Theta) = L_e(x \rightarrow \Theta) + \int_{\Omega_x} f_r(\Psi \leftrightarrow \Theta) \cdot L(x \leftarrow \Psi) \cdot \cos \theta_x \cdot d\omega_\Psi$$

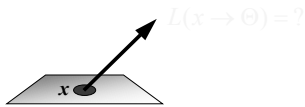
- Area integration (over polygons from set A)

$$L(x \rightarrow \Theta) = L_e(x \rightarrow \Theta) + \int_A f_r(\Psi \leftrightarrow \Theta) \cdot L(y \rightarrow -\Psi) \cdot \frac{\cos \theta_x \cdot \cos \theta_y}{r_{xy}^2} V(x, y) \cdot dA_y$$

© Kavita Bala, Computer Science, Cornell University

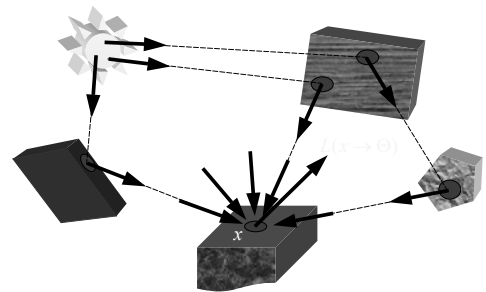
Radiance evaluation

- Fundamental problem in GI algorithms
 - Evaluate radiance at a given surface point in a given direction
 - Invariance defines radiance everywhere else



© Kavita Bala, Computer Science, Cornell University

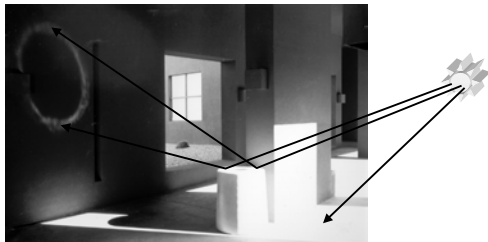
Radiance evaluation



... find paths between sources and surfaces to be shaded

© Kavita Bala, Computer Science, Cornell University

Hard to find paths

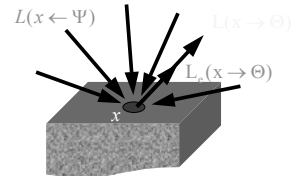


© Kavita Bala, Computer Science, Cornell University

Why Monte Carlo?

- Radiance is hard to evaluate

$$L(x \rightarrow \Theta) = L_e(x \rightarrow \Theta) + \int_{\Omega_x} f_r(\Psi \leftrightarrow \Theta) \cdot L(x \leftarrow \Psi) \cdot \cos(\Psi, n_x) \cdot d\omega_\Psi$$

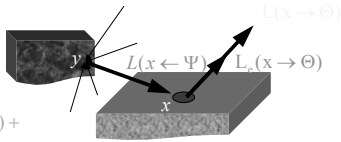


- Sample many paths: integrate over all incoming directions

© Kavita Bala, Computer Science, Cornell University

Why Monte Carlo?

$$L(x \rightarrow \Theta) = L_c(x \rightarrow \Theta) + \int_{\Omega_x} f_r(\Psi \leftrightarrow \Theta) \cdot L(x \leftarrow \Psi) \cdot \cos(\Psi, n_x) \cdot d\omega_\Psi$$



$$L(x \rightarrow \Theta) = L_c(x \rightarrow \Theta) +$$

$$\int_{\Omega_x} L_c(y \rightarrow -\Psi) f_r(\Psi \leftrightarrow \Theta) \cdot \cos(\Psi, n_x) \cdot d\omega_\Psi +$$

$$\int_{\Omega_x} \int_{\Omega_x} f_r(\Psi' \rightarrow -\Psi) \cos(\Psi', n_x) L(y \leftarrow \Psi') d\omega_{\Psi'} f_r(\Psi \leftrightarrow \Theta) \cdot \cos(\Psi, n_x) \cdot d\omega_\Psi$$

© Kavita Bala, Computer Science, Cornell University

Why Monte Carlo?

$$L(x \rightarrow \Theta) = L_c(x \rightarrow \Theta) +$$

$$\int_{\Omega_x} L_c(y \rightarrow -\Psi) f_r(\Psi \leftrightarrow \Theta) \cdot \cos(\Psi, n_x) \cdot d\omega_\Psi +$$

.....

- Analytical integration is difficult
- Therefore, need numerical techniques

© Kavita Bala, Computer Science, Cornell University

Monte Carlo Integration

- Numerical tool to evaluate integrals
- Use sampling
- Stochastic errors
- Unbiased
 - on average, we get the right answer!

© Kavita Bala, Computer Science, Cornell University

Probability

- Random variable x
- Possible outcomes: $x_1, x_2, x_3, \dots, x_n$
 - each with probability: $p_1, p_2, p_3, \dots, p_n$
- E.g. 'average die': 2,3,3,4,4,5
 - outcomes: $x_1 = 2, x_2 = 3, x_3 = 4, x_4 = 5$
 - probabilities:

$$p_1 = 1/6, p_2 = 1/3, p_3 = 1/3, p_4 = 1/6$$

© Kavita Bala, Computer Science, Cornell University

Expected value

- Expected value = average value

$$E[x] = \sum_{i=1}^n x_i p_i$$

- E.g. die:

$$E[x] = 2 \cdot \frac{1}{6} + 3 \cdot \frac{1}{3} + 4 \cdot \frac{1}{3} + 5 \cdot \frac{1}{6} = 3.5$$

© Kavita Bala, Computer Science, Cornell University

Variance

- Expected 'distance' to expected value

$$\sigma^2[x] = E[(x - E[x])^2]$$

- E.g. die:

$$\sigma^2[x] = (2-3.5)^2 \cdot \frac{1}{6} + (3-3.5)^2 \cdot \frac{1}{3} + (4-3.5)^2 \cdot \frac{1}{3} + (5-3.5)^2 \cdot \frac{1}{6}$$

$$= 0.916$$

- Property: $\sigma^2[x] = E[x^2] - E[x]^2$

© Kavita Bala, Computer Science, Cornell University

Continuous random variable

- Random variable $x \in [a, b]$
- Probability density function (pdf) $p(x)$
- Probability that variable has value x : $p(x)dx$

$$\int_a^b p(x)dx = 1$$

- Cumulative distribution function (CDF)
 - CDF is non-decreasing, positive

$$\Pr(x \leq y) = CDF(y) = \int_{-\infty}^y p(x)dx$$

© Kavita Bala, Computer Science, Cornell University

Continuous random variable

- Expected value: $E[x] = \int_a^b xp(x)dx$

$$E[g(x)] = \int_a^b g(x)p(x)dx$$

- Variance:

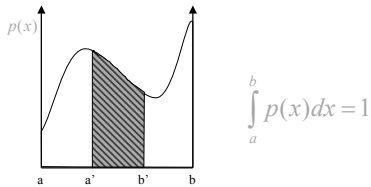
$$\sigma^2[x] = \int_a^b (x - E[x])^2 p(x)dx$$

$$\sigma^2[g(x)] = \int_a^b (g(x) - E[g(x)])^2 p(x)dx$$

- Deviation: $\sigma[x], \sigma[g(x)]$

© Kavita Bala, Computer Science, Cornell University

Continuous random variable



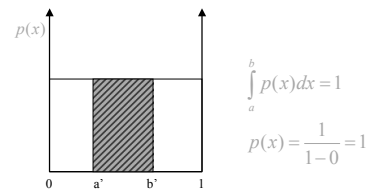
$$\int_a^b p(x)dx = 1$$

$$\Pr(x \leq y) = CDF(y) = \int_{-\infty}^y p(x)dx$$

$$\begin{aligned} \text{Probability that } x \text{ belongs to } [a', b'] &= \Pr(x \leq b') - \Pr(x \leq a') \\ &= \int_{-\infty}^{b'} p(x)dx - \int_{-\infty}^{a'} p(x)dx = \int_{a'}^{b'} p(x)dx \end{aligned}$$

© Kavita Bala, Computer Science, Cornell University

Uniform distribution



$$\int_a^b p(x)dx = 1$$

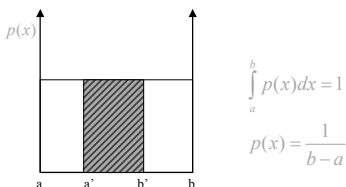
$$p(x) = \frac{1}{1-0} = 1$$

$$\Pr(x \in [a', b']) = \int_{a'}^{b'} 1dx = (b' - a')$$

$$\Pr(x \leq y) = CDF(y) = \int_{-\infty}^y p(x)dx = y$$

© Kavita Bala, Computer Science, Cornell University

Uniform distribution



$$\int_a^b p(x)dx = 1$$

$$p(x) = \frac{1}{b-a}$$

$$\text{Probability that } x \text{ belongs to } [a', b'] = \int_{a'}^{b'} \frac{1}{(b-a)} dx = \frac{(b'-a')}{(b-a)}$$

$$\Pr(x \leq y) = CDF(y) = \int_{-\infty}^y p(x)dx = \frac{(y-a)}{(b-a)}$$

© Kavita Bala, Computer Science, Cornell University

More than one sample

- Consider the weighted sum of N samples
- Expected value $E[\frac{1}{N}(x^1 + x^2 + x^3 + \dots + x^N)] = E[x]$

- Variance $\sigma^2[\frac{1}{N}(x^1 + x^2 + x^3 + \dots + x^N)] = \frac{1}{N}\sigma^2[x]$

- Deviation $\sigma[\frac{1}{N}(x^1 + x^2 + x^3 + \dots + x^N)] = \frac{1}{\sqrt{N}}\sigma[x]$

© Kavita Bala, Computer Science, Cornell University

More than one sample

- Consider the weighted sum of N samples

$$g(x) = \frac{1}{N}(f(x_1) + f(x_2) + f(x_3) + \dots + f(x_N))$$

- Expected value

$$E[g(x)] = E\left[\frac{1}{N} \sum_{i=1}^N f(x_i)\right] = E[f(x)]$$

- Variance

$$\sigma^2[g(x)] = \sigma^2\left[\frac{1}{N} \sum_{i=1}^N f(x_i)\right] = \frac{1}{N} \sigma^2[f(x)]$$

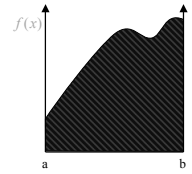
- Deviation $\sigma[g(x)] = \frac{1}{\sqrt{N}} \sigma[f(x)]$

© Kavita Bala, Computer Science, Cornell University

Numerical Integration

- A one-dimensional integral:

$$I = \int_a^b f(x) dx$$

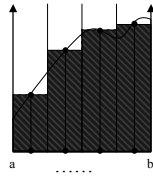


© Kavita Bala, Computer Science, Cornell University

Deterministic Integration

- Quadrature rules:

$$I = \int_a^b f(x) dx \approx \sum_{i=1}^N w_i f(x_i)$$



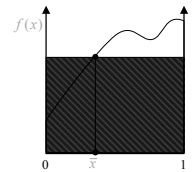
© Kavita Bala, Computer Science, Cornell University

Monte Carlo Integration

- Primary estimator:

$$I = \int_a^b f(x) dx$$

$$I_{prim} = f(\bar{x})$$



© Kavita Bala, Computer Science, Cornell University

Monte Carlo Integration

- Primary estimator:

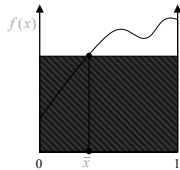
$$I = \int_a^b f(x) dx$$

$$I_{prim} = f(\bar{x})$$

$$E(I_{prim}) = \int_0^1 f(x) p(x) dx = \int_0^1 f(x) 1 dx = I$$

Unbiased estimator!

© Kavita Bala, Computer Science, Cornell University



Monte Carlo Integration

- Primary estimator:

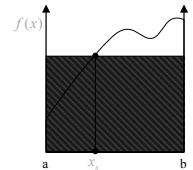
$$I = \int_a^b f(x) dx$$

$$I_{prim} = f(x_s)(b-a)$$

$$E(I_{prim}) = \int_a^b f(x)(b-a)p(x) dx = \int_a^b f(x)(b-a) \frac{1}{(b-a)} dx = I$$

Unbiased estimator!

© Kavita Bala, Computer Science, Cornell University



Monte Carlo Integration: Error

Variance of the estimator → a measure of the stochastic error

$$\sigma_{prim}^2 = \int_a^b \left[\frac{f(x)}{p(x)} - I \right]^2 p(x) dx$$

© Kavita Bala, Computer Science, Cornell University

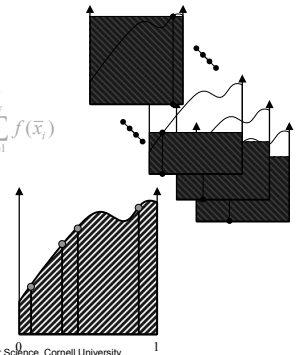
More samples

Secondary estimator

Generate N random samples x_i

$$\text{Estimator: } \langle I \rangle = I_{sec} = \frac{1}{N} \sum_{i=1}^N f(\bar{x}_i)$$

$$\text{Variance } \sigma_{sec}^2 = \sigma_{prim}^2 / N$$



© Kavita Bala, Computer Science, Cornell University

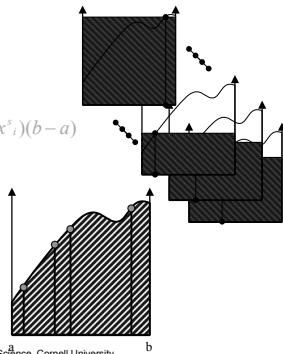
More samples

Secondary estimator

Generate N random samples x_i

$$\text{Estimator: } \langle I \rangle = I_{sec} = \frac{1}{N} \sum_{i=1}^N f(x_i)(b-a)$$

$$\text{Variance } \sigma_{sec}^2 = \sigma_{prim}^2 / N$$



© Kavita Bala, Computer Science, Cornell University

Monte Carlo Integration

• Expected value of estimator

$$\begin{aligned} E[\langle I \rangle] &= E\left[\frac{1}{N} \sum_{i=1}^N \frac{f(x_i)}{p(x_i)}\right] = \frac{1}{N} \int \left(\sum_{i=1}^N \frac{f(x_i)}{p(x_i)}\right) p(x) dx \\ &= \frac{1}{N} \sum_{i=1}^N \int \left(\frac{f(x)}{p(x)}\right) p(x) dx \\ &= \frac{N}{N} \int f(x) dx = I \end{aligned}$$

– on ‘average’ get right result: **unbiased**

• Standard deviation σ is a measure of the stochastic error

$$\sigma^2 = \frac{1}{N} \int \left[\frac{f(x)}{p(x)} - I \right]^2 p(x) dx$$

© Kavita Bala, Computer Science, Cornell University

Convergence Rates

- RMS error converges at a rate of $O\left(\frac{1}{\sqrt{N}}\right)$
- Unbiased
- Chebychev's inequality

$$\Pr\left[|F - E(F)| \geq \sqrt{\frac{1}{\delta}} \sigma\right] \leq \delta$$

$$\Pr\left[|I_{estimator} - I| \geq \frac{1}{\sqrt{N}} \sqrt{\frac{1}{\delta}} \sigma\right] \leq \delta$$

- Strong law of large numbers

$$\Pr\left[\lim_{N \rightarrow \infty} \frac{1}{N} \sum_{i=1}^N \frac{f(x_i)}{p(x_i)} = I\right] = 1$$

© Kavita Bala, Computer Science, Cornell University

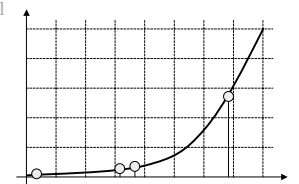
MC Integration - Example

– Integral $I = \int_0^1 5x^4 dx = 1$

– Uniform sampling

– Samples :

$x_1 = .86$	$\langle I \rangle = 2.74$
$x_2 = .41$	$\langle I \rangle = 1.44$
$x_3 = .02$	$\langle I \rangle = 0.96$
$x_4 = .38$	$\langle I \rangle = 0.75$



© Kavita Bala, Computer Science, Cornell University

MC Integration - Example

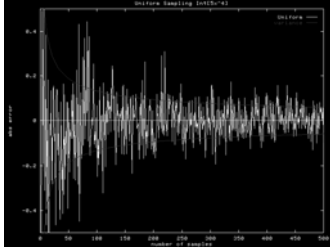
- Integral

$$I = \int_0^1 5x^4 dx = 1$$

- Stochastic error

- Variance

- What is it?

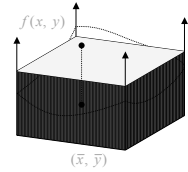


© Kavita Bala, Computer Science, Cornell University

MC Integration: 2D

- Primary estimator:

$$\bar{I}_{prim} = \frac{f(\bar{x}, \bar{y})}{p(\bar{x}, \bar{y})}$$

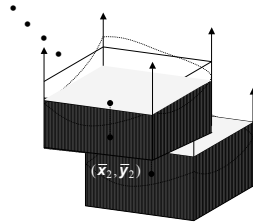


© Kavita Bala, Computer Science, Cornell University

MC Integration: 2D

- Secondary estimator:

$$I_{sec} = \frac{1}{N} \sum_{i=1}^N \frac{f(\bar{x}_i, \bar{y}_i)}{p(\bar{x}_i, \bar{y}_i)}$$



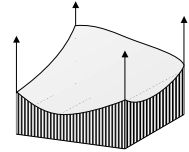
© Kavita Bala, Computer Science, Cornell University

Monte Carlo Integration - 2D

- MC Integration works well for higher dimensions
- Unlike quadrature

$$I = \int_a^b \int_c^d f(x, y) dx dy$$

$$\langle I \rangle = \frac{1}{N} \sum_{i=1}^N \frac{f(x_i, y_i)}{p(x_i, y_i)}$$



© Kavita Bala, Computer Science, Cornell University

MC Advantages

- Convergence rate of $O\left(\frac{1}{\sqrt{N}}\right)$
- Simple
 - Sampling
 - Point evaluation
 - Can use black boxes
- General
 - Works for high dimensions
 - Deals with discontinuities, crazy functions,...

© Kavita Bala, Computer Science, Cornell University

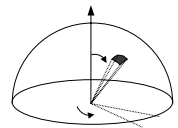
MC Integration - 2D example

- Integration over hemisphere:

$$I = \int_{\Omega} f(\Theta) d\omega_{\Theta}$$

$$= \int_0^{2\pi} \int_0^{\pi/2} f(\varphi, \theta) \sin \theta d\theta d\varphi$$

$$\langle I \rangle = \frac{1}{N} \sum_{i=1}^N \frac{f(\varphi_i, \theta_i) \sin \theta_i}{p(\varphi_i, \theta_i)}$$



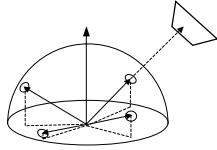
© Kavita Bala, Computer Science, Cornell University

Hemisphere Integration example

Irradiance due to light source:

$$I = \int_{\Omega} L_{source} \cos \theta d\omega_{\odot}$$

$$= \int_0^{2\pi} \int_0^{\pi/2} L_{source} \cos \theta \sin \theta d\theta d\phi$$



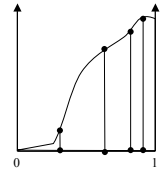
$$p(\omega_i) = \frac{\cos \theta \sin \theta}{\pi}$$

$$\langle I \rangle = \frac{1}{N} \sum_{i=1}^N \frac{L_{source}(\omega_i) \cos \theta \sin \theta}{p(\omega_i)} = \frac{\pi}{N} \sum_{i=1}^N L_{source}(\omega_i)$$

© Kavita Bala, Computer Science, Cornell University

Importance Sampling

- Better use of samples by taking more samples in 'important' regions, i.e. where the function is large



© Kavita Bala, Computer Science, Cornell University

MC integration - Non-Uniform

- Some parts of the integration domain have higher importance
- Generate samples according to density function $p(x)$

$$\langle I \rangle = \frac{1}{N} \sum_{i=1}^N \frac{f(x_i)}{p(x_i)}$$

- Estimator?

- What is optimal $p(x)$? $p(x) \approx f(x) / \int f(x) dx$

© Kavita Bala, Computer Science, Cornell University

MC integration - Non-Uniform

- Generate samples according to density function $p(x)$

$$p(x) \approx f(x) / \int f(x) dx$$

- Why? $I_{estimator} = \frac{1}{N} \sum \frac{f(x)}{p(x)} = \frac{1}{N} \sum \frac{f(x)}{f(x)/I} = \frac{1}{N} \sum I = I$

$$\sigma^2 = \frac{1}{N} \int \left[\frac{f(x)}{p(x)} - I \right]^2 p(x) dx$$

- But.....

$$= \frac{1}{N} \int \left[\frac{f(x)}{f(x)/I} - I \right]^2 p(x) dx = 0$$

© Kavita Bala, Computer Science, Cornell University

Example

- Function: $I = \int_0^1 x dx = 8$ $f(x) = x$

$$\sigma^2 = \frac{1}{N} \int_a^b \left[\frac{f(x)}{p(x)} - I \right]^2 p(x) dx$$

$$p(x) = \frac{x}{8}, \sigma^2 = 0 \quad I_{estimator} = I = 8$$

$$p(x) = \frac{1}{4}, \sigma^2 = \frac{1}{N} \int_0^4 \left[\frac{x}{1/4} - 8 \right]^2 \frac{1}{4} dx = 21.3 / N$$

$$p(x) = \frac{x+2}{16}, \sigma^2 = \frac{1}{N} \int_0^4 \left[\frac{x}{(x+2)/16} - 8 \right]^2 \frac{x+2}{16} dx = 6.3 / N$$

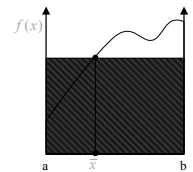
© Kavita Bala, Computer Science, Cornell University

Monte Carlo Integration

Primary estimator:

$$I = \int_a^b f(x) dx$$

$$I_{prim} = f(\bar{x})(b-a)$$



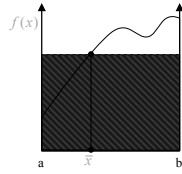
© Kavita Bala, Computer Science, Cornell University

Monte Carlo Integration

Primary estimator:

$$I = \int_a^b f(x) dx$$

$$I_{prim} = f(\bar{x})(b-a)$$



$$E(I_{prim}) = \int_a^b f(x)(b-a)p(x) dx = \int_a^b f(x)(b-a) \frac{1}{(b-a)} dx = I$$

Unbiased estimator!

© Kavita Bala, Computer Science, Cornell University

Monte Carlo Integration: Error

Variance of the estimator:

$$\sigma_{prim}^2 = \int_a^b \left[\frac{f(x)}{p(x)} - I \right]^2 p(x) dx$$

→ a measure of the stochastic error

© Kavita Bala, Computer Science, Cornell University

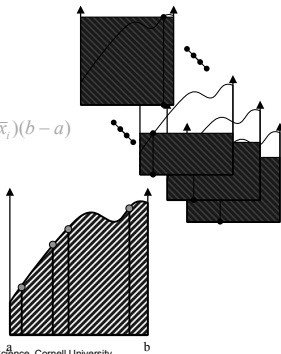
More samples

Secondary estimator

Generate N random samples x_i ,

$$\text{Estimator: } \langle I \rangle = I_{sec} = \frac{1}{N} \sum_{i=1}^N f(\bar{x}_i)(b-a)$$

$$\text{Variance } \sigma_{sec}^2 = \sigma_{prim}^2 / N$$



© Kavita Bala, Computer Science, Cornell University