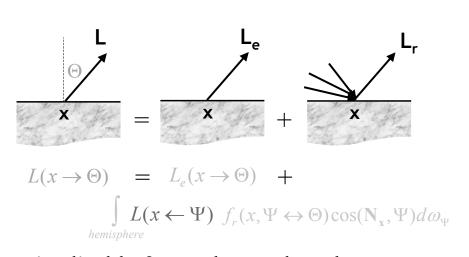
# Lecture 6: Monte Carlo Integration Chapter 3 in Advanced GI

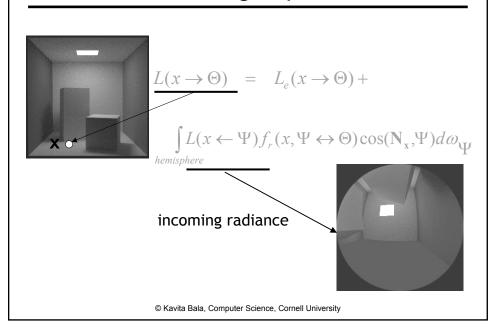
Fall 2004
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Computer Science
Cornell University

## **Rendering Equation**



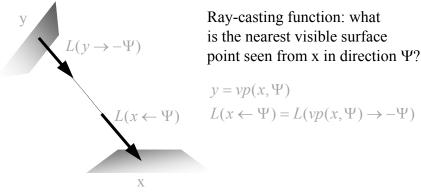
• Applicable for each wavelength

## Rendering Equation



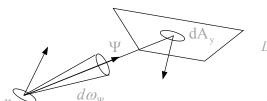
#### RE: Area Formulation

$$L(x \to \Theta) = L_e(x \to \Theta) + \int_{\Omega_x} f_r(\Psi \leftrightarrow \Theta) \cdot L(x \leftarrow \Psi) \cdot \cos \theta_x \cdot d\omega_{\Psi}$$



#### Rendering Equation

$$L(x \to \Theta) = L_e(x \to \Theta) + \int_{\Omega_x} f_r(\Psi \leftrightarrow \Theta) \cdot L(x \leftarrow \Psi) \cdot \cos \theta_x \cdot d\omega_{\Psi}$$



$$y = vp(x, \Upsilon)$$

$$L(x \leftarrow \Psi) = L(vp(x, \Psi) \rightarrow -\Psi)$$

$$d\omega_{\Psi} = \frac{dA_y \cos \theta_y}{r_{xy}^2}$$

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## Rendering Equation: visible surfaces

$$\begin{split} L(x \to \Theta) &= L_e(x \to \Theta) + \int\limits_{\Omega_x} f_r(\Psi \leftrightarrow \Theta) \cdot L(x \leftarrow \Psi) \cdot \cos \theta_x \cdot d\omega_\Psi \\ &\text{Coordinate transform} \quad \bigg | \end{split}$$

$$L(x \to \Theta) = L_e(x \to \Theta) + \int_{\substack{y \text{ on} \\ \text{all surfaces}}}^{y \text{ on}} f_r(\Psi \leftrightarrow \Theta) \cdot L(y \to -\Psi) \cos \theta_x \cdot \frac{\cos \theta_y}{r_{xy}^2} \cdot dA_y$$

Integration domain = visible surface points y

· Integration domain extended to ALL surface points by including visibility function

#### Rendering Equation: all surfaces

$$L(x \to \Theta) = L_e(\ldots) + \int_A f_r(\ldots) \cdot L(y \to -\Psi) \cdot \frac{\cos \theta_x \cdot \cos \theta_y}{r_{xy}^2} \cdot V(x,y) \, dA_y$$
 
$$\otimes \text{ Kavita Bala, Computer Science, Cornell University}$$

#### Two forms of the RE

Hemisphere integration

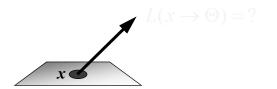
$$L(x \to \Theta) = L_e(x \to \Theta) + \int_{\Omega_x} f_r(\Psi \leftrightarrow \Theta) \cdot L(x \leftarrow \Psi) \cdot \cos \theta_x \cdot d\omega_{\Psi}$$

Area integration (over polygons from set A)

$$L(x \to \Theta) = L_e(x \to \Theta) + \int_A f_r(\Psi \leftrightarrow \Theta) \cdot L(y \to -\Psi) \cdot \frac{\cos \theta_x \cdot \cos \theta_y}{r_{xy}^2} \cdot V(x, y) \cdot dA_y$$

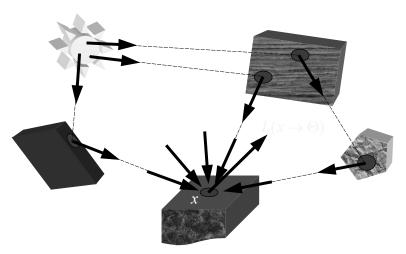
#### Radiance evaluation

- Fundamental problem in GI algorithms
  - Evaluate radiance at a given surface point in a given direction
  - Invariance defines radiance everywhere else



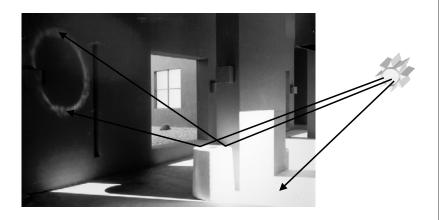
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#### Radiance evaluation



... find paths between sources and surfaces to be shaded

# Hard to find paths



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# Why Monte Carlo?

· Radiance is hard to evaluate

$$\underline{L(x \to \Theta)} = \underline{L_e(x \to \Theta)} + \int_{\Omega_x} f_r(\Psi \leftrightarrow \Theta) \cdot \underline{L(x \leftarrow \Psi)} \cdot \cos(\Psi, n_x) \cdot d\omega_{\Psi}$$

$$\underline{L(x \leftarrow \Psi)}$$

$$\underline{L(x \to \Theta)}$$

Sample many paths: integrate over all incoming directions

#### Why Monte Carlo?

$$L(x \to \Theta) = L_e(x \to \Theta) + \int_{\Omega_x} f_r(\Psi \leftrightarrow \Theta) \cdot L(x \leftarrow \Psi) \cdot \cos(\Psi, n_x) \cdot d\omega_{\Psi}$$

$$L(x \to \Theta) = L_e(x \to \Theta) + \\ \int\limits_{\Omega_x} L_e(y \to -\Psi) f_r(\Psi \leftrightarrow \Theta) \cdot \cos(\Psi, n_x) \cdot d\omega_\Psi + \\ \int\limits_{\Omega_x} f_r(\Psi' \to -\Psi) \cos(\Psi', n_y') L(y \leftarrow \Psi') d\omega_{\psi'} f_r(\Psi \leftrightarrow \Theta) \cdot \cos(\Psi, n_x) \cdot d\omega_\Psi \\ \otimes \text{ Kavita Bala, Computer Science, Cornell University}$$

## Why Monte Carlo?

$$\begin{split} L(x \to \Theta) &= L_e(x \to \Theta) + \\ &\int\limits_{\Omega_x} L_e(y \to -\Psi) f_r(\Psi \leftrightarrow \Theta) \cdot \cos(\Psi, n_x) \cdot d\omega_\Psi + \end{split}$$

- Analytical integration is difficult
- Therefore, need numerical techniques

#### Monte Carlo Integration

- Numerical tool to evaluate integrals
- Use sampling
- Stochastic errors
- Unbiased
  - on average, we get the right answer!

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# Probability

- Random variable x
- Possible outcomes:  $x_1, x_2, x_3, ..., x_n$ 
  - each with probability:  $p_1, p_2, p_3, ..., p_n$
- E.g. 'average die': 2,3,3,4,4,5
  - **outcomes:**  $x_1 = 2, x_2 = 3, x_3 = 4, x_3 = 5$
  - probabilities:

$$p_1 = 1/6, p_2 = 1/3, p_3 = 1/3, p_3 = 1/6$$

#### Expected value

• Expected value = average value

$$E[x] = \sum_{i=1}^{n} x_i p_i$$

• E.g. die:

$$E[x] = 2 \cdot \frac{1}{6} + 3 \cdot \frac{1}{3} + 4 \cdot \frac{1}{3} + 5 \cdot \frac{1}{6} = 3.5$$

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#### Variance

• Expected 'distance' to expected value

$$\sigma^{2}[x] = E[(x - E[x])^{2}]$$

• E.g. die:

$$\sigma^{2}[x] = (2 - 3.5)^{2} \cdot \frac{1}{6} + (3 - 3.5)^{2} \cdot \frac{1}{3} + (4 - 3.5)^{2} \cdot \frac{1}{3} + (5 - 3.5)^{2} \cdot \frac{1}{6}$$
$$= 0.916$$

• **Property:**  $\sigma^2[x] = E[x^2] - E[x]^2$ 

#### Continuous random variable

- Random variable  $x \in [a,b]$
- Probability density function (pdf) p(x)
- Probability that variable has value x: p(x)dx

$$\int_{a}^{b} p(x)dx = 1$$

- Cumulative distribution function (CDF)
  - CDF is non-decreasing, positive

$$Pr(x \le y) = CDF(y) = \int_{-\infty}^{y} p(x)dx$$

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#### Continuous random variable

• Expected value: 
$$E[x] = \int_{a}^{b} xp(x)dx$$

$$E[g(x)] = \int_{a}^{b} g(x)p(x)dx$$

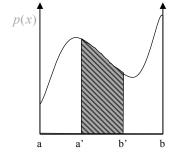
$$\sigma^{2}[x] = \int_{a}^{b} (x - E[x])^{2} p(x) dx$$

$$\sigma^{2}[x] = \int_{a}^{b} (x - E[x])^{2} p(x) dx$$

$$\sigma^{2}[g(x)] = \int_{a}^{b} (g(x) - E[g(x)])^{2} p(x) dx$$

• Deviation:  $\sigma[x], \sigma[g(x)]$ 

#### Continuous random variable



$$\int_{a}^{b} p(x)dx = 1$$

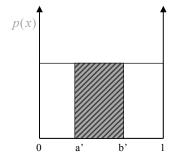
$$Pr(x \le y) = CDF(y) = \int_{-\infty}^{y} p(x)dx$$

Probability that x belongs to [a',b'] =  $Pr(x \le b') - Pr(x \le a')$ 

$$= \int_{-\infty}^{b'} p(x)dx - \int_{-\infty}^{a'} p(x)dx = \int_{a'}^{b'} p(x)dx$$

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#### Uniform distribution



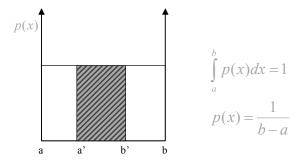
$$\int_{a}^{b} p(x)dx = 1$$

$$p(x) = \frac{1}{1 - 0} = 1$$

$$\Pr(x \in [a', b']) = \int_{a'}^{b'} 1 dx = (b' - a')$$

$$Pr(x \le y) = CDF(y) = \int_{-\infty}^{y} p(x)dx = y$$

#### Uniform distribution



Probability that x belongs to [a',b'] = 
$$\int_{a'}^{b'} \frac{1}{(b-a)} dx = \frac{(b'-a')}{(b-a)}$$

$$Pr(x \le y) = CDF(y) = \int_{-\infty}^{y} p(x)dx = \frac{(y-a)}{(b-a)}$$

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#### More than one sample

- Consider the weighted sum of N samples
- Expected value  $E[\frac{1}{N}(x^1 + x^2 + x^3 + ... x^N)] = E[x]$

• Variance 
$$\sigma^2[\frac{1}{N}(x^1 + x^2 + x^3 + ... x^N)] = \frac{1}{N}\sigma^2[x]$$

• **Deviation** 
$$\sigma[\frac{1}{N}(x^1 + x^2 + x^3 + ... x^N)] = \frac{1}{\sqrt{N}}\sigma[x]$$

#### More than one sample

Consider the weighted sum of N samples

$$g(x) = \frac{1}{N}(f(x_1) + f(x_2) + f(x_3) + \dots + f(x_N))$$

Expected value

$$E[g(x)] = E[\frac{1}{N} \sum_{i=1}^{N} f(x_i)] = E[f(x)]$$

Variance

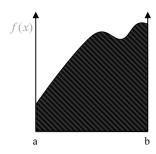
$$\sigma^{2}[g(x)] = \sigma^{2}[\frac{1}{N}\sum_{i}^{N}f(x_{i})] = \frac{1}{N}\sigma^{2}[f(x)]$$

**Deviation** 
$$\sigma[g(x)] = \frac{1}{\sqrt{N}}\sigma[f(x)]$$

# **Numerical Integration**

• A one-dimensional integral:

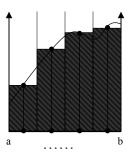
$$I = \int_{a}^{b} f(x) dx$$



# **Deterministic Integration**

Quadrature rules:

$$I = \int_{a}^{b} f(x)dx$$
$$\approx \sum_{i=1}^{N} w_{i} f(x_{i})$$



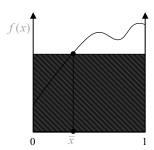
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# Monte Carlo Integration

#### Primary estimator:

$$I = \int_{a}^{b} f(x) dx$$

$$I_{prim} = f(\overline{x})$$

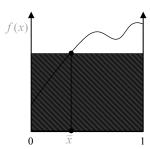


#### Monte Carlo Integration

#### Primary estimator:

$$I = \int_{a}^{b} f(x) dx$$

$$I_{prim} = f(\overline{x})$$



$$E(I_{prim}) = \int_{0}^{1} f(x)p(x)dx = \int_{0}^{1} f(x)1dx = I$$

#### Unbiased estimator!

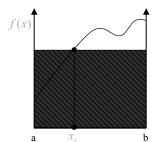
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# Monte Carlo Integration

#### Primary estimator:

$$I = \int_{a}^{b} f(x) dx$$

$$I_{prim} = f(x_s)(b-a)$$



$$E(I_{prim}) = \int_{a}^{b} f(x)(b-a)p(x)dx = \int_{a}^{b} f(x)(b-a)\frac{1}{(b-a)}dx = I$$

#### Unbiased estimator!

# Monte Carlo Integration: Error

# Variance of the estimator $\rightarrow$ a measure of the stochastic error

$$\sigma_{prim}^2 = \int_a^b \left[ \frac{f(x)}{p(x)} - I \right]^2 p(x) dx$$

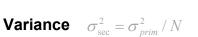
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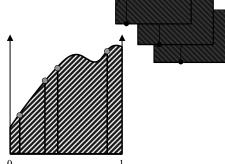
## More samples

#### **Secondary estimator**

Generate N random samples  $\boldsymbol{x}_i$ 

$$\langle I \rangle = I_{\text{sec}} = \frac{1}{N} \sum_{i=1}^{N} f(\bar{x}_i)$$





#### More samples

#### Secondary estimator

Generate N random samples x<sub>i</sub>

Estimator:  $\langle I \rangle = I_{\text{sec}} = \frac{1}{N} \sum_{i=1}^{N} f(x^{s_i})(b-a)$ 

**Variance** 
$$\sigma_{\rm sec}^2 = \sigma_{\it prim}^2 / N$$

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# Monte Carlo Integration

· Expected value of estimator

$$E[\langle I \rangle] = E[\frac{1}{N} \sum_{i}^{N} \frac{f(x_{i})}{p(x_{i})}] = \frac{1}{N} \int (\sum_{i}^{N} \frac{f(x_{i})}{p(x_{i})}) p(x) dx$$
$$= \frac{1}{N} \sum_{i}^{N} \int (\frac{f(x)}{p(x)}) p(x) dx$$
$$= \frac{N}{N} \int f(x) dx = I$$

- on 'average' get right result: unbiased
- Standard deviation  $\sigma$  is a measure of the stochastic error  $\sigma^2 = \frac{1}{N} \int \left[ \frac{f(x)}{p(x)} - I \right]^2 p(x) dx$

#### Convergence Rates

- RMS error converges at a rate of O(  $\frac{1}{\sqrt{N}}$  )
- Unbiased
- Chebychev's inequality

$$\begin{split} & \Pr \Bigg[ \left| F - E(F) \right| \geq \sqrt{\frac{1}{\delta}} \sigma \, \Bigg] \leq \delta \\ & \Pr \Bigg[ \left| I_{\textit{estimator}} - I \right| \geq \frac{1}{\sqrt{N}} \sqrt{\frac{1}{\delta}} \sigma \, \Bigg] \leq \delta \end{split}$$

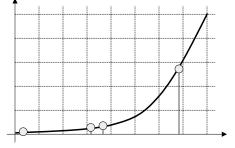
Strong law of large numbers

$$\Pr\left[\lim_{N\to\infty}\frac{1}{N}\sum_{i}^{N}\frac{f(x_{i})}{p(x_{i})}=I\right]=1$$

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# MC Integration - Example

- Integral
- $I = \int 5x^4 dx = 1$
- Uniform sampling
- Samples:



$$x_1 = .86$$

$$x_1 = .86$$
  *= 2.74*

$$x_2 = .41$$
  *= 1.44*

$$< I > = 1.44$$

$$x_3 = .02$$

$$< I > = 0.96$$

$$x_4 = .38$$

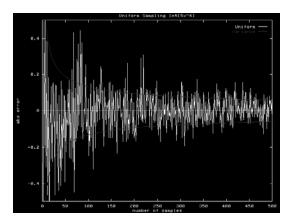
$$< I > = 0.75$$

# MC Integration - Example

- Integral

$$I = \int_{0}^{1} 5x^{4} dx = 1$$

- Stochastic error
- Variance
  - What is it?

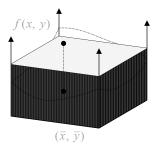


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# MC Integration: 2D

• Primary estimator:

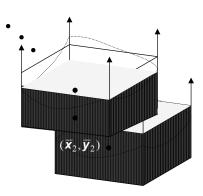
$$\bar{I}_{prim} = \frac{f(\bar{x}, \bar{y})}{p(\bar{x}, \bar{y})}$$



# MC Integration: 2D

Secondary estimator:

$$I_{\text{sec}} = \frac{1}{N} \sum_{i=1}^{N} \frac{f(\overline{x}_i, \overline{y}_i)}{p(\overline{x}_i, \overline{y}_i)}$$



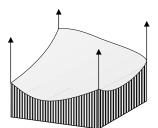
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# Monte Carlo Integration - 2D

- MC Integration works well for higher dimensions
- Unlike quadrature

$$I = \int_{a}^{b} \int_{c}^{d} f(x, y) dx dy$$

$$\langle I \rangle = \frac{1}{N} \sum_{i=1}^{N} \frac{f(x_i, y_i)}{p(x_i, y_i)}$$



#### MC Advantages

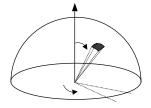
- Convergence rate of O( $\frac{1}{\sqrt{N}}$ )
- Simple
  - Sampling
  - Point evaluation
  - Can use black boxes
- General
  - Works for high dimensions
  - Deals with discontinuities, crazy functions,...

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## MC Integration - 2D example

Integration over hemisphere:

$$I = \int_{\Omega} f(\Theta) d\omega_{\Theta}$$
$$= \int_{0}^{2\pi\pi/2} \int_{0}^{\pi/2} f(\varphi, \theta) \sin \theta d\theta d\varphi$$

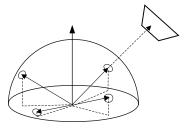


$$\langle I \rangle = \frac{1}{N} \sum_{i=1}^{N} \frac{f(\varphi_i, \theta_i) \sin \theta}{p(\varphi_i, \theta_i)}$$

#### Hemisphere Integration example

#### Irradiance due to light source:

$$\begin{split} I &= \int\limits_{\Omega} L_{source} \cos\theta d\omega_{\Theta} \\ &= \int\limits_{0}^{2\pi\pi/2} \int\limits_{0}^{2\pi\pi/2} L_{source} \cos\theta \sin\theta d\theta d\phi \end{split}$$



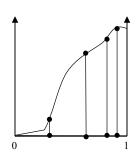
$$p(\omega_i) = \frac{\cos\theta\sin\theta}{\pi}$$

$$\langle I \rangle = \frac{1}{N} \sum_{i=1}^{N} \frac{L_{source}(\omega_i) \cos \theta \sin \theta}{p(\omega_i)} = \frac{\pi}{N} \sum_{i=1}^{N} L_{source}(\omega_i)$$

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# Importance Sampling

 Better use of samples by taking more samples in 'important' regions, i.e. where the function is large



## MC integration - Non-Uniform

- Some parts of the integration domain have higher importance
- Generate samples according to density function p(x)

 $\langle I \rangle = \frac{1}{N} \sum_{i=1}^{N} \frac{f(x_i)}{p(x_i)}$ 

- Estimator?
- What is optimal p(x)?  $p(x) \approx f(x) / \int f(x) dx$

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## MC integration - Non-Uniform

• Generate samples according to density function p(x)  $p(x) \approx f(x) / \int f(x) dx$ 

• Why? 
$$I_{estimator} = \frac{1}{N} \sum \frac{f(x)}{p(x)} = \frac{1}{N} \sum \frac{f(x)}{f(x)/I} = \frac{1}{N} \sum I = I$$

• But....
$$\sigma^{2} = \frac{1}{N} \int_{a}^{b} \left[ \frac{f(x)}{p(x)} - I \right]^{2} p(x) dx$$

$$= \frac{1}{N} \int_{a}^{b} \left[ \frac{f(x)}{f(x)/I} - I \right]^{2} p(x) dx = 0$$

#### Example

$$I = \int_{0}^{4} x dx = 8$$

• Function: 
$$I = \int_{0}^{4} x dx = 8$$
 
$$f(x) = x$$
 
$$\sigma^{2} = \frac{1}{N} \int_{a}^{b} \left[ \frac{f(x)}{p(x)} - I \right]^{2} p(x) dx$$

$$p(x) = \frac{x}{8}, \sigma^2 = 0$$
  $I_{estimator} = I = 8$ 

$$p(x) = \frac{1}{4}, \sigma^2 = \frac{1}{N} \int_0^4 \left[ \frac{x}{1/4} - 8 \right]^2 \frac{1}{4} dx = 21.3/N$$

$$p(x) = \frac{x+2}{16}, \sigma^2 = \frac{1}{N} \int_0^4 \left[ \frac{x}{(x+2)/16} - 8 \right]^2 \frac{x+2}{16} dx = 6.3/N$$

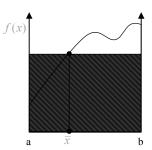
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# Monte Carlo Integration

#### Primary estimator:

$$I = \int_{a}^{b} f(x) dx$$

$$I_{prim} = f(\overline{x})(b-a)$$

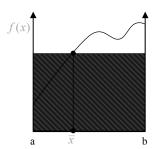


#### Monte Carlo Integration

#### Primary estimator:

$$I = \int_{a}^{b} f(x) dx$$

$$I_{prim} = f(\overline{x})(b - a)$$



$$E(I_{prim}) = \int_{a}^{b} f(x)(b-a)p(x)dx = \int_{a}^{b} f(x)(b-a)\frac{1}{(b-a)}dx = I$$

#### Unbiased estimator!

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# Monte Carlo Integration: Error

#### Variance of the estimator:

$$\sigma_{prim}^2 = \int_a^b \left[ \frac{f(x)}{p(x)} - I \right]^2 p(x) dx$$

→ a measure of the stochastic error

# More samples

#### **Secondary estimator**

Generate N random samples  $x_i$ 

Estimator:  $\langle I \rangle = I_{\text{sec}} = \frac{1}{N} \sum_{i=1}^{N} f(\overline{x}_i)(b-a)$ 

Variance 
$$\sigma_{\rm sec}^2 = \sigma_{\it prim}^2 / N$$