# **Expected Utility Theory**

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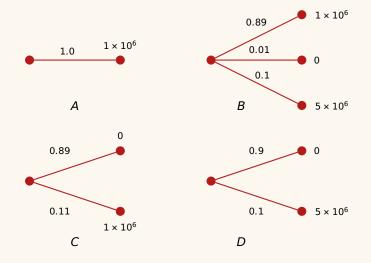
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# **Powerball Payouts**

Prize Level	Payout	Odds
Match 5 + PB	Jackpot	1 in 292,201,338
Match 5	\$1,000,000	1 in 1,688,054
Match 4 + PB	\$50,000	1 in 913,129
Match 4	\$100	1 in 36,525
Match 3 + PB	\$100	1 in 14,494
Match 3	\$7	1 in 580
Match 2 + PB	\$7	1 in 701
Match 1 + PB	\$4	1 in 92
Match 0 + PB	\$4	1 in 38

Source: http://www.powerball.net

The Jackpot is currently estimated at \$90 mil. — \$68 mil. cash value

Payout	Probability	
$68 \times 10^{6}$	$3.4223 \times 10^{-9}$	
$1 \times 10^{6}$	$5.99501 \times 10^{-7}$	
50,000	$1.09514 \times 10^{-6}$	
100	9.63726 × 10 <sup>-5</sup>	
7	$3.15067 \times 10^{-3}$	
4	$3.71854 \times 10^{-2}$	

Expected winnings are \$1.0603. A ticket costs \$2, so the net expected gain is -\$0.9397.

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For a net expected payoff of 0 the cash payout must be \$342.6 mil.

#### An Important Fact!

**Deductions for Gambling Losses** 

Playing the lottery is classed as gambling as far as the Internal Revenue Service (IRS) is concerned, which means that you are entitled to a tax deduction on any losses incurred. To file these deductions, you will need to keep an accurate record of your wins and losses, as well as any evidence of them, such as the tickets you bought. You must itemize the deductions on the tax form 1040, obtainable from the IRS website. The losses you deduct cannot exceed your income from all forms of gambling, including but not limited to horse racing, casinos, and raffles.

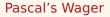
Source https://www.powerball.net/taxes

Am I 17 million times better off winning \$68 million than I am winning \$4?.

Pascal's Wager



	God Exists	There is No God
Succeed in Believing	Eternal Life	Finite, Deluded Life
Remain an Atheist	Oh, Hell!	What I now presume





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	God Exists	There is No God
Succeed in Believing	8	x
Remain an Atheist	У	Z

At what odds should you choose to remain an atheist?

#### The Saint Petersburg Paradox

Here is a game:

- Flip a fair coin until tails comes up.
- If the first T appears at flip N, you are paid \$2<sup>N</sup>.

How much would you be willing to pay to play this game?



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$$\mathsf{E}\left\{2^{\tilde{N}}\right\} = \frac{1}{2} \cdot 2 + \frac{1}{4} \cdot 4 + \dots + \frac{1}{2^{n}} \cdot 2^{n} + \dots = +\infty$$

# History of EU

- ▶ Daniel Bernoulli's resolution of the St. Petersburg paradox was to introduce expected utility. He observed that although  $\sum_n 2^{-n} 2^n$  is unbounded,  $\sum_n 2^{-n} \log 2^n = 2 \log 2$ . If you were to pay *w* for the wager, your expected utility is  $\sum_n 2^{-n} \log(2^n w)$ . This is a good bet for w <\$1.81.
- It came under criticism in the 1930s.
  - People evaluate gambles by looking at the mean, the variance, and other statistics. (Hicks 1931)
  - The utility function whose expectation is being taken is "cardinal". (Tintner 1942)

So the late 40's and early 50's brought forth the zoo of functions.

# History of EU

- von Neumann and Morgenstern (1947) provided the first axiomatization of EU preferences.
- Much of the acceptance of EU in the early 50s came from the normative force of the axioms. But there has been pushback against the descriptive validity of EU since the late 40s. Milton Friedman and Leonard Savage wrote a famous and famously bad paper in 1948 that tried to reconcile the fact that people buy both insurance and lottery tickets with EU. Few of the major players in 1950s decision theory thought that EU was empirically valid.

 Additive separability across states - independence axiom, sure thing principle

Stochastic monotonicity

Representation properties e.g. risk aversion

When does a preference relation have an EU representation?

X Outcomes of lotteries (finite for now)

 $\mathscr{L} = \{p_1, x_1; \dots, p_N, x_N\}$  A simple lottery with probabilities  $p_1, \dots, p_N$  and prizes  $x_1, \dots, x_N$ .

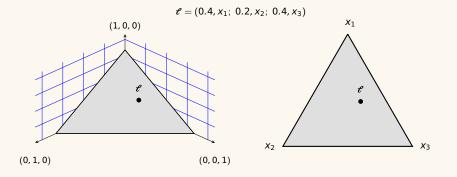
△ is the set  $\{(p_1, ..., p_N) : p_n \ge 0 \& \sum_n p_n = 1\}$ .

# Characterization of Preferences Notation

Kreps provides three levels of organization

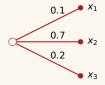
- A fixed & finite prize space. Lotteries are elements in Δ.
   Useful for drawing pictures.
- An infinite set of possible prizes, but lotteries are probability distributions with **finite support**, that is, each lottery assigns probability 1 to some finite set of prizes. This is what Kreps means by a "simple lottery". The set of such lotteries is a mixture space. Kreps calls this P<sub>S</sub> but since P is overloaded in this class I will call it *S*.
- Lotteries with countable and continuum support. Kreps doesn't name it and neither will I.

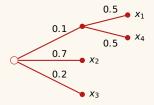
# **Three Alternatives**



#### A representation of $\Delta$ and $\mathscr{L}$ .

# Simple & Compound Lotteries

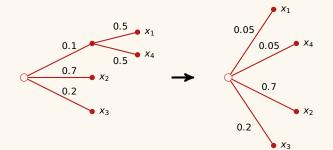




A Simple Lottery

A Compound Lottery

# **Reduction of Compound Lotteries**



#### **Mixture Spaces**

Definition. A mixture space is a non-empty set  $\mathcal{M}$  together with an operation

$$[0,1] \times \mathcal{M} \times \mathcal{M} \to \mathcal{M}$$
$$(\lambda, \ell, m) \to \ell \lambda m$$

s.t.

A.1  $\ell lm = \ell$ , A.2  $\ell \lambda m = m(1-\lambda)\ell$ , A.3  $(\ell \lambda m)\mu m = \ell(\lambda \mu)m$ .

A function  $U : \mathcal{M} \to \mathbb{R}$  is mixture-preserving if for all  $\lambda, \ell, m$ 

$$U(\ell \lambda m) = \lambda U(\ell) + (1-\lambda)U(m)$$

# Some Properties of Mixture Spaces

$$\ell 0m = m \text{Proof. } \ell 0m \stackrel{\text{A2}}{=} m 1\ell \stackrel{\text{A1}}{=} m.$$

$$\mathcal{C} \lambda \mathcal{C} = \mathcal{C}$$
Proof.  $\ell \lambda \mathcal{C} \stackrel{\text{A1}}{=} (\ell 1 \ell) \lambda \mathcal{C} \stackrel{\text{A2}}{=} (\ell 0 \ell) \lambda \mathcal{C} \stackrel{\text{A3}}{=} \ell 0 \mathcal{C} = \ell$ 

(
$$\ell \lambda m$$
) $\mu(\ell \nu m) = \ell(\lambda \nu + (1 - \lambda)\nu)m$   
Proof. See Fishburn (1982).

#### **Examples of Mixture Spaces**

• A convex set is a mixture space with  $\ell \lambda m = \lambda \ell + (1 - \lambda)m$ .

► 
$$\mathcal{M} = \{\ell, m, n\}$$
. For all  $\lambda, m\lambda m = m$  and  $n\lambda n = n$ ,  
 $\ell 0m = m1\ell = n0m = m1n = m$ ,  
 $\ell 0n = n1\ell = m0n = n1m = n$ , and all other mixtures equal  
 $\ell$ .

• 
$$\mathcal{M} = \{\ell, m, n\}$$
. For all  $\lambda \in (0, 1)$ ,  $\ell \lambda m = m \lambda \ell = m$ ,  
 $m \lambda n = n \lambda m = n$ ,  $n \lambda \ell = \ell \lambda n = \ell$ , and the 0,1 mixtures are  
chosen according to A.1 and A.2.

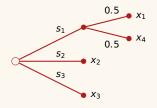
Here are two properties of convex sets not shared by all mixture spaces:

$$\blacktriangleright (\ell \beta m) \alpha n = \ell \alpha \beta (q \alpha (1 - \beta) (1 - \alpha \beta)^{-1} n) \text{ (associativity)}$$

• for  $\alpha \neq 0$ ,  $\ell \alpha n = m \alpha n$  implies that  $\ell = m$ . (determinacy)

# Why Mixture Spaces?

Why not just reduce compound lotteries to simple lotteries and embed them in a convex set as the picture on slide 14 suggests?



Reduce this!

 $s_1$ ,  $s_2$  and  $s_3$  are states of nature without given probabilities.

#### The Mixture Space Representation Theorem

**Theorem**. Let  $\succ$  be a binary relation on a mixture space  $\mathcal{M}$ . Assume:

- ► (Preference) > is a preference relation,
- (Archimedean) For all  $\ell$ , m,  $n \in \mathcal{M}$  such that  $\ell \succ m \succ n$ , there are  $0 < \alpha, \beta < 1$  such that

$$\ell \alpha n \succ m \succ \ell \beta n$$
,

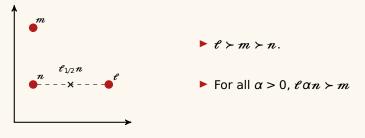
► (Independence) For all  $\ell$ ,  $m, n \in M$  and  $0 < \alpha \le 1$ , if  $\ell \succ m$  then

$$\ell \alpha n \succ m \alpha n$$
.

Then  $\succ$  has a mixture-preserving representation:  $m \succ n$  iff U(m) > U(n), and  $U(m\alpha n) = \alpha U(m) + (1 - \alpha)U(n)$ . Furthermore, if V is another mixture-preserving representation, for some  $\alpha > 0$  and  $\beta$ ,

$$V(m) = \alpha U(m) + \beta.$$

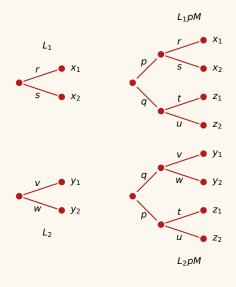
# Archimedean Axiom



The Lexicographic Order

#### Independence Axiom

The usual justification goes as follows:



Independence:  $L_1 \succ L_2$  iff  $L_1 pM \succ L_2 pM$ .

Suppose  $L_1$  is preferred to  $L_2$ . Now imagine flipping a coin and getting  $L_1$  on H and a default lottery M on T. How should it compare to getting  $L_2$  on H and the same default on T.

This sounds normatively plausible. Descriptively the reduction of compound lotteries is questionable. The Allais paradox (to come) provides a counterexample.

#### Proof of the Representation Theorem

Here are a bunch of facts that can be quickly derived:

• If  $m \succ n$  and  $0 \ge \alpha < \beta \ge 1$  then  $m\beta n \succ m\alpha n$ .

This is a monotonicity property.

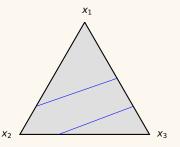
► If  $\ell \succcurlyeq m \succcurlyeq n$  and  $\ell \succ n$ , then for exactly one  $0 \le \alpha \le 1$  $m \sim \ell \alpha n$ .

Among other things, this suggests that indifference curves are not thick.

▶ If  $\ell \sim m$  then  $\ell \alpha n \sim m \alpha n$ 

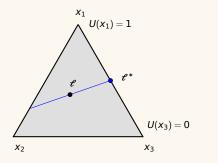
This says two things:

- If  $\ell \sim m$  then for all  $0 \le \alpha \le 1$  $\ell \alpha m \sim m \alpha m = m$ , Indifference curves are lines.
- Indifference curves are parallel.



#### Proof of the Representation Theorem

Idea. Suppose  $x_1$  is the best mixture and  $x_3$  is the worst.



 $U(\mathcal{C})$  solves  $\mathcal{C} \sim \mathcal{C}^*$  and  $\mathcal{C}^* = x_1 U(\mathcal{C}) x_3$ 

This idea extends to the case where there is no best and worst mixture. See Kreps p. 55.

#### Application to von Neumann–Morgenstern EU

**Theorem.** Let  $\mathscr{L}$  denote the set of lotteries on a finite outcome space X and let  $\succ$  be a preference order satisfying axioms A.1.–3. Then there exists a  $u : X \rightarrow \mathbb{R}$  such that

$$U(p) = E_p u \equiv \sum_n p_n u(x_n)$$

represents  $\succ$ . Furthermore,  $v : X \rightarrow \mathbb{R}$  similarly represents  $\succ$  iff  $v = \alpha u + \beta$  with  $\alpha > 0$ .

**Proof.**  $\mathscr{D}$  with the convex combination operation is a mixture space. The mixture space representation theorem gives a representation  $U: \mathscr{C} \to \mathbb{R}$  such that  $U(p\gamma q) = \gamma U(p) + (1 - \gamma)U(q)$ . Arguing by induction on the cardinality of X proves that  $U(p) = \sum_{n} p_n u(x_n)$ .

# Cardinal Utility?

A relational system  $\Re = \langle X, R_1, \dots, R_K \rangle$  is a set X of objects together with K relations. These relations may be binary, ternary, etc.

If, for example,  $\mathscr{R} = (X, R_1, R_2, R_3)$  is a relational system, and if  $R_1$  and  $R_3$  are binary while  $R_2$  is ternary, we say that the type of  $\mathscr{R}$  is 2, 3, 2.

A function *F* is a numerical representation of the relational system  $\mathscr{R}$  iff there is a real relational system (on the object set of real numbers) of the same type,  $\langle \mathbb{R}, S_1, \ldots, S_K \rangle$  and a function  $F : X \to \mathbb{R}$  such that for all k,  $(x_1, \ldots, x_{m_k}) \in R_k$  iff  $((F(x_1), \ldots, F(x_{m_k})) \in S_k$ .

# Cardinal Utility?

It is often said that the "utility" u is a cardinal measure of utility on X, because any other "utility" v is related to u by a positive affine transformation.

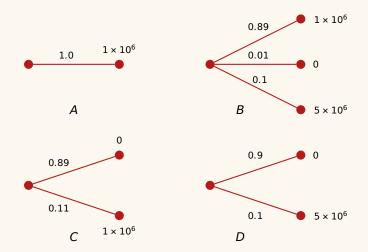
u(w) - u(x) > u(y) - u(z) iff v(w) - v(x) > v(y) - v(z).

Since relations between utility differences are invariant under all "utilities" that appear in vn-M representations of  $\succ$ , it must be they are significant, that there is an implicit quaternary relationship R' on X, to wit,  $(w, x, y, z) \in R'$  iff the decision maker prefers w over x more than she prefers y over z. It is claimed that this is meaningful.

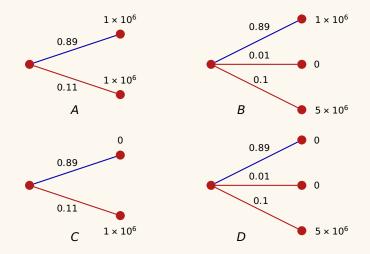
If this were true, then we would have derived the quaternary relationship "more better than" from the binary relationship "better than".

What's wrong with this?

# The Allais Paradox



# The Allais Paradox



#### Calling Out BS

Consider the following mixture space:

- O is a finite set of outcomes.
- >  $X_0$  is the set of sure prizes.
- ►  $X_n$  consists of all **binary** lotteries on  $X_{n-1}$ ,  $(\lambda, x; (1-\lambda), y)$ where  $0 \le \lambda \le 1$  and  $x, y \in X_{n-1}$ .
- ►  $X = \cup_n X_n$ .

Identify the elements (1, x; 0, y) and x. Then  $X_{n-1} \subset X_n$ . This lets us define on X a mixture operator: For  $x \in X_m$  and  $y \in X_n$  define

$$x\lambda y = (\lambda, x; (1-\lambda), y) \in X_{\max\{m,n\}+1}.$$

If  $\succ$  on X satisfies A.1–3 then it has a mixture-preserving representation.

#### Calling Out BS

For distinct outcomes  $o_1, o_2, o_3, o_1\lambda(o_2\gamma o_3) \in X_2$ .

$$U(o_1\lambda(o_2\gamma o_3)) = \lambda U(o_1) + (1-\lambda)U(o_2\gamma o_3)$$
  
=  $\lambda U(o_1) + (1-\lambda)(\gamma U(o_2) + (1-\gamma)U(o_3))$   
=  $\lambda U(o_1) + (1-\lambda)\gamma U(o_2) + (1-\lambda)(1-\gamma)U(o_3)$ 

but although  $(\lambda, o_1; (1-\gamma)(\gamma, o_2; (1-\gamma), o_3)$  exists in  $X_2$ , the lottery  $(\lambda, o_1; (1-\lambda)\gamma, o_2; (1-\lambda)(1-\gamma)o_3)$  does not exist in X.

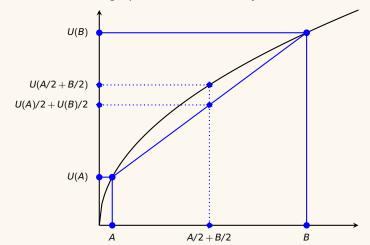
**Definition**. A decision maker is risk averse if for any gamble  $\ell = (p_1, x_1; ..., p_N, x_N), (1, \sum_n p_n x_n) \geq \ell$ .

In expected utility terms,

 $U(\mathsf{E}\{X\}) \ge \mathsf{E}\{U(X)\}$ 

This will be true for all gambles iff *U* is concave.

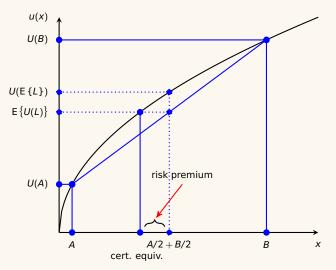
#### **Concave Functions**



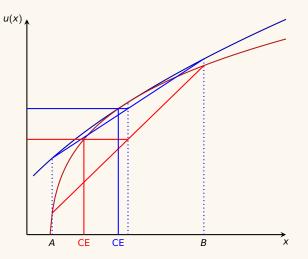
A concave function's graph is on or above any of its secant lines.

# **Concave Functions**

The lottery L = (1/2, A; 1/2, B).



# Curvature and Risk Aversion



### Measuring Curvature

Curvatures are measured by coefficients of risk aversion.

The coefficient of absolute risk aversion is

$$\rho_A(x) = -\frac{u''(x)}{u'(x)} = \frac{d}{dx} \log u'(x)$$

The coefficient of relative risk aversion is

$$\rho_R(x) = -\frac{xu''(x)}{u'(x)} = \frac{d}{d\log x}\log u'(x)$$

#### **Constant Risk Aversion**

Proposition A utility function u has Constant Absolute Risk Aversion  $\kappa > 0$  iff

$$u(x) = -e^{-\kappa x}$$

Proposition A utility function *u* has Constant Relative Risk Aversion  $\gamma > 0$ . Iff

$$u(x)=\frac{1}{1-\gamma}x^{1-\gamma}$$

or, for  $\gamma = 1$ ,

 $u(x) = \log x.$ 

The coefficients are preserved by positive affine transformations.

#### **Risk Aversion**

Homework problem: Show that if u is CARA and F is a gamble whose payoff is normally distributed with mean  $\mu$  and variance  $\sigma^2$ , then

$$CE(F) = \mu - \kappa \sigma^2$$

**Proposition.** If *u* is any increasing, concave and  $C^2$  payoff function, and  $F_{\sigma}$  is a family of gambles with fixed mean  $\mu$  and variance  $\sigma^2$ , then if  $\sigma^2$  is sufficiently small,

$$CE(F_{\sigma}) \approx \mu - \rho_A(\mu)\sigma^2$$

What does it mean to say that individual *Y* is at least as risk averse than is individual *Z*?

- For any gamble F,  $CE^{Y}(F) \ge CE^{Z}(F)$ ,
- For all x,  $\rho_A^Y(x) > \rho^Z(x)$ .
- ▶ There is a concave function  $g : \mathbb{R} \to \mathbb{R}$  such that  $u^Y = g \circ u^Z$ .

These are all equivalent.

# An Application

An individual has an initial wealth w and may lose 1 unit with probability p. She can buy insurance. To insure x dollars of the loss costs  $q \cdot x$ . Under what circumstances will she buy insurance, and when she buys, how much?

If she buys x dollars of insurance, her expected utility is

$$U(x) = (1-p)u(w-qx) + pu(w-qx-1+x).$$

Actuarially fair insurance requires q = p.

#### **Unfair Insurance**

Suppose 
$$q > p$$
. Will she choose  $x = 1$ ?  
 $U'(1) = (p(1-q) - q(1-p))u'(w-q) = (p-q)u'(w-q) < 0$ 

so NO!

Suppose q = p. Will she choose x = 1?

$$U'(1) = (p(1-q) - q(1-p))u'(w-q) = (p-q)u'(w-q) = 0$$

so YES! For any x, E  $\{\tilde{w}\} = w - p$ . For x < 1,

$$CE(x) < E\{\tilde{w}\} = w - q = CE(1)$$

so x = 1 has a higher certainty equivalence than x < 1, hence is optimal.

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