

Machine Learning for Data Science (CS4786)

Lecture 20

Finish HMM, Inference in Graphical Models

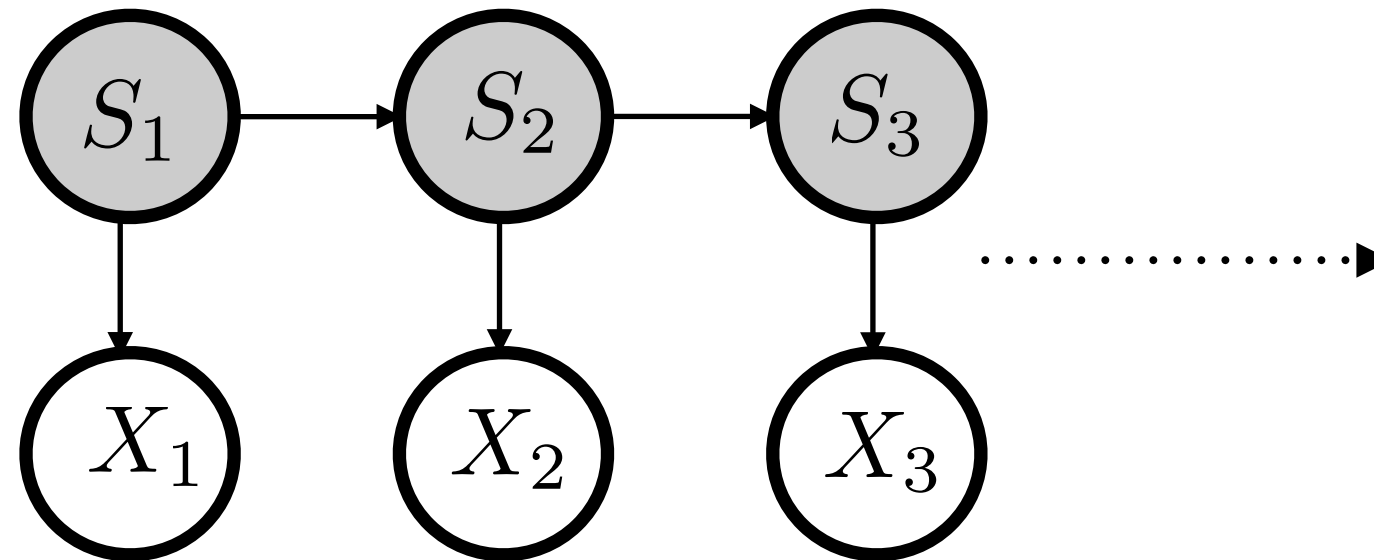
Course Webpage :

<http://www.cs.cornell.edu/Courses/cs4786/2016fa/>

ANNOUNCEMENT

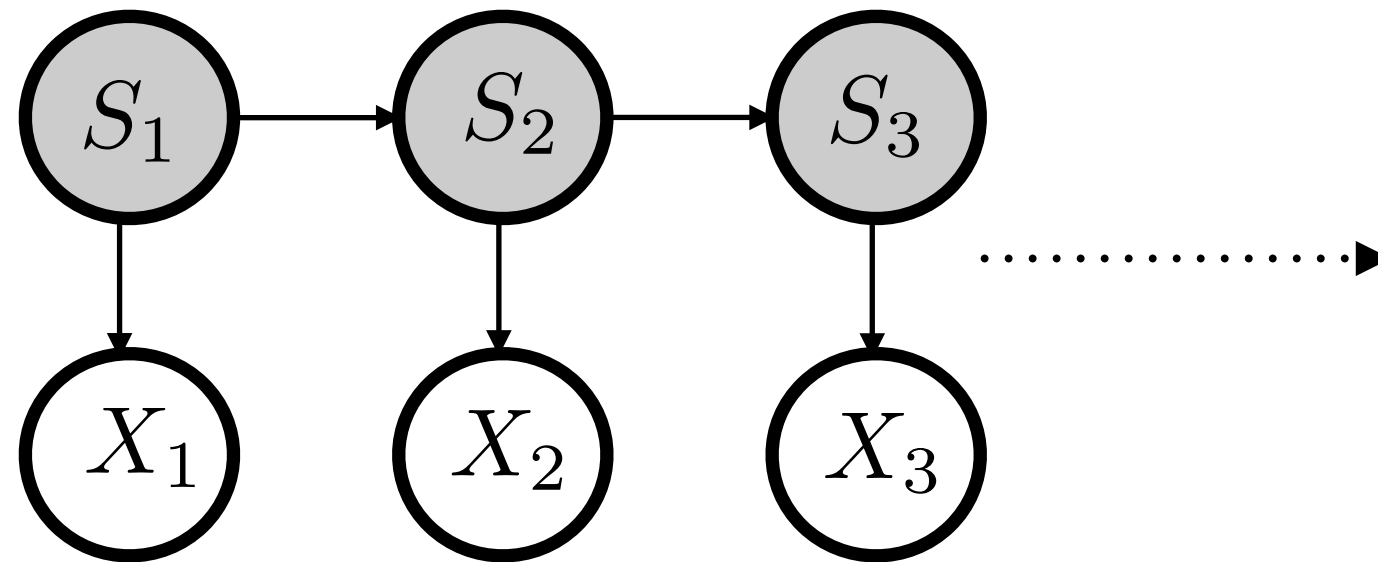
- Good job on competition I, report due today
- No lecture Tuesday, Nov 8th!
- Next Thursday Nov 10th, guest lecture by Prof. Kilian Weinberger on TSNE

HIDDEN MARKOV MODEL (HMM)



Parameters: Transition probability matrix T
Emission probability matrix E

INFERENCE IN HMM

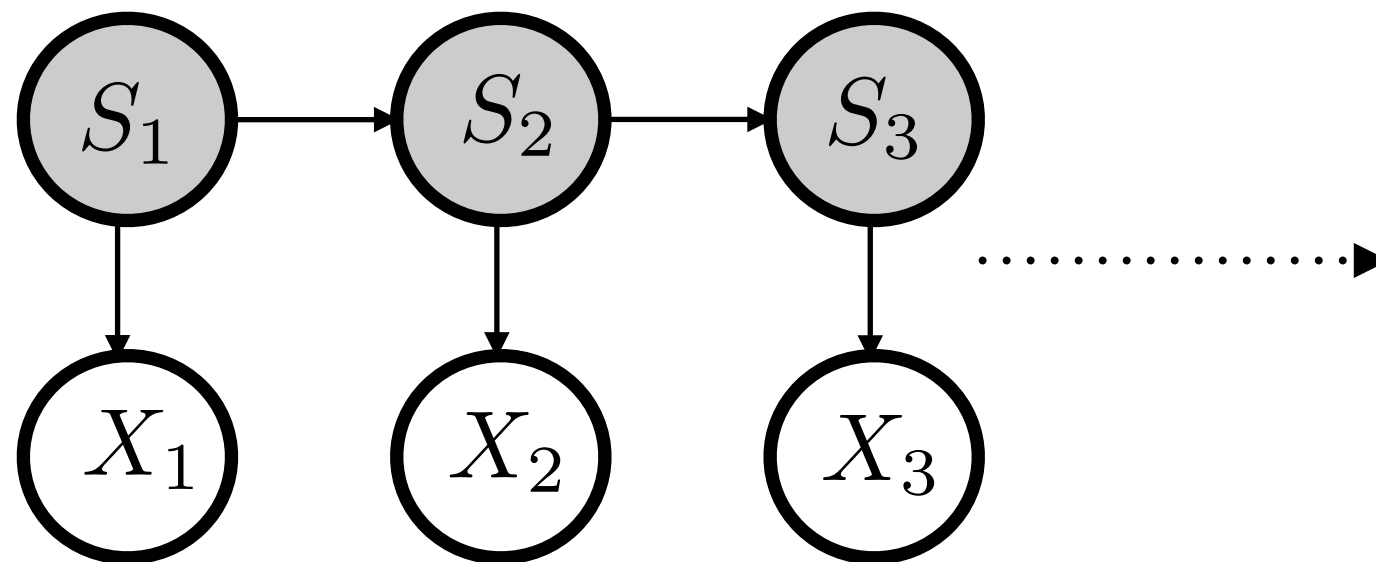


$$\text{message}_{S_{t-1} \mapsto S_t}(k) = P(S_t = k, X_1, \dots, X_{t-1})$$

$$\text{message}_{S_{t+1} \mapsto S_t}(k) = P(X_n, \dots, X_{t+1} | S_t = k)$$

$$P(S_t = k | X_1, \dots, X_n) \propto \text{message}_{S_{t-1} \mapsto S_t}(k) \times \text{message}_{S_{t+1} \mapsto S_t}(k) \times P(X_t | S_t = k)$$

INFERENCE IN HMM



$$\text{message}_{S_{t-1} \mapsto S_t}(k) = P(S_t = k, X_1, \dots, X_{t-1})$$

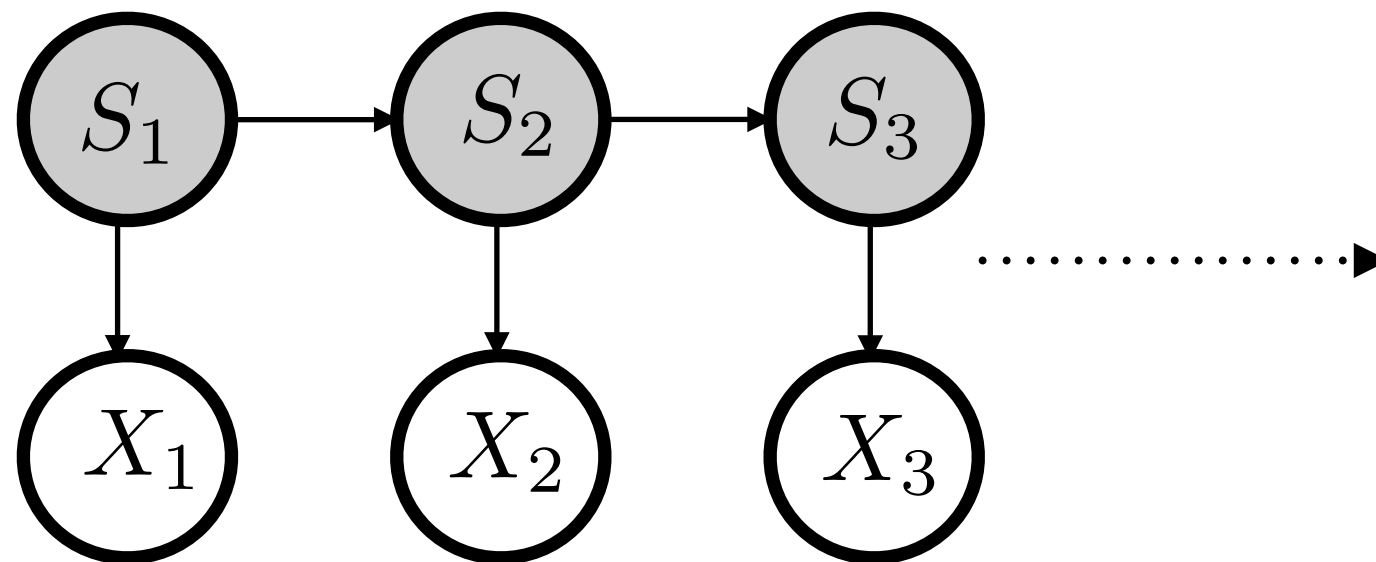
$$\text{message}_{S_{t+1} \mapsto S_t}(k) = P(X_n, \dots, X_{t+1} | S_t = k)$$

Forward:

$$P(X_1, \dots, X_{t-1}, S_t = k) = \sum_{j=1}^K P(S_t = k | S_{t-1} = j) P(X_{t-1} | S_{t-1} = j) P(X_1, \dots, X_{t-2}, S_{t-1} = j)$$

$$\text{message}_{S_{t-1} \mapsto S_t}(k) = \sum_{j=1}^K P(S_t = k | S_{t-1} = j) P(X_{t-1} | S_{t-1} = j) \text{message}_{S_{t-2} \mapsto S_{t-1}}(j)$$

INFERENCE IN HMM



$$\text{message}_{S_{t-1} \mapsto S_t}(k) = P(S_t = k, X_1, \dots, X_{t-1})$$

$$\text{message}_{S_{t+1} \mapsto S_t}(k) = P(X_n, \dots, X_{t+1} | S_t = k)$$

Backward:

$$P(X_n, \dots, X_{t+1} | S_t = k) = \sum_{j=1}^K P(X_n, \dots, X_{t+2} | S_{t+1} = j) P(X_{t+1} | S_{t+1} = j) P(S_{t+1} = j | S_t = k)$$

$$\text{message}_{S_{t+1} \mapsto S_t}(k) = \sum_{j=1}^K \text{message}_{S_{t+2} \mapsto S_{t+1}}(j) P(X_{t+1} | S_{t+1} = j) P(S_{t+1} = j | S_t = k)$$

LEARNING PARAMETERS FOR HMM

- Now that we have algorithm for inference, what about learning
- Given observations, how do we estimate parameters for HMM?
Three guesses ...

EM FOR HMM (BAUM WELCH)

- EM algorithm of course, for HMM its referred to as Baum Welch algorithm
- Initialize Transition and Emission probability tables arbitrarily
- For $i = 1$ to convergence:

E-step For every state variable $t \in \{1, \dots, n\}$,
Use forward-backward algorithm to compute probabilities of latent variables given observation

M-step Optimize weighted log likelihood as usual:

$$\theta^{(i)} = \arg \max_{\theta \in \Theta} \sum_{S_{1,\dots,n}} P(S_{1,\dots,n} | X_{1,\dots,n}, \theta^{(i-1)}) \log P(X_{1,\dots,n}, S_{1,\dots,n} | \theta)$$

LETS SIMPLIFY M-STEP

$$\begin{aligned}\log P(X_{1,\dots,n}, S_{1,\dots,n}|\theta) &= \log \left(\prod_{t=1}^n P(X_t|S_t, \theta) \prod_{t=1}^n P(S_t|S_{t-1}, \theta) \right) \\ &= \sum_{t=1}^n \log P(X_t|S_t, \theta) + \sum_{t=1}^n \log P(S_t|S_{t-1}, \theta)\end{aligned}$$

Hence,

$$\begin{aligned}&\sum_{S_{1,\dots,n}} P(S_{1,\dots,n}|X_{1,\dots,n}, \theta^{(i-1)}) \log P(X_{1,\dots,n}, S_{1,\dots,n}|\theta) \\ &= \sum_{t=1}^n \sum_{s_t=1}^K P(S_t = s_t|X_{1,\dots,n}, \theta^{i-1}) \log P(X_t|S_t = s_t, \theta) \\ &\quad + \sum_{t=1}^n \sum_{s_t, s_{t-1}=1}^K P(S_t = s_t, S_{t-1} = s_{t-1}|X_{1,\dots,n}, \theta^{i-1}) \log P(S_t|S_{t-1}, \theta)\end{aligned}$$

E-STEP

- Only need to compute $P(S_t = s_t | X_{1,\dots,n}, \theta^{i-1})$ and $P(S_t = s_t, S_{t-1} = s_{t-1} | X_{1,\dots,n}, \theta^{i-1})$ using forward-backward
- First term is immediate

$$P(S_t = s_t | X_{1,\dots,n}, \theta^{i-1}) \propto m_{S_{t-1} \mapsto S_t}(s_t) \cdot m_{S_{t+1} \mapsto S_t}(s_t) \cdot E^{(i-1)}[s_t, X_t]$$

- For second term,

$$\begin{aligned} & P(S_t = s_t, S_{t-1} = s_{t-1} | X_{1,\dots,n}, \theta^{i-1}) \\ & \propto m_{S_{t-1} \mapsto S_t}(s_t) T^{(i-1)}[s_{t-1}, s_t] P(S_{t-1} = s_{t-1} | X_{1,\dots,n}, \theta^{i-1}) \\ & \propto m_{S_{t-1} \mapsto S_t}(s_t) T^{(i-1)}[s_{t-1}, s_t] m_{S_{t-2} \mapsto S_{t-1}}(s_{t-1}) m_{S_t \mapsto S_{t-1}}(s_{t-1}) E^{(i-1)}[s_{t-1}, X_{t-1}] \end{aligned}$$

Why?

E-STEP

$$\begin{aligned} & P(S_t = s_t, S_{t-1} = s_{t-1} | X_{1,\dots,n}, \theta^{i-1}) \\ &= P(S_t = s_t, | S_{t-1} = s_{t-1}, X_{1,\dots,n}, \theta^{i-1}) P(S_{t-1} = s_{t-1} | X_{1,\dots,n}, \theta^{i-1}) \\ &= P(S_t = s_t, | S_{t-1} = s_{t-1}, X_{t,\dots,n}, \theta^{i-1}) P(S_{t-1} = s_{t-1} | X_{1,\dots,n}, \theta^{i-1}) \\ &\propto P(X_{t,\dots,n} | S_t = s_t, S_{t-1} = s_{t-1}, \theta^{i-1}) \\ &\quad P(S_t = s_t | S_{t-1} = s_{t-1}, \theta^{i-1}) P(S_{t-1} = s_{t-1} | X_{1,\dots,n}, \theta^{i-1}) \\ &\propto P(X_{t,\dots,n} | S_t = s_t, \theta^{i-1}) \\ &\quad T^{(i-1)}[s_{t-1}, s_t] P(S_{t-1} = s_{t-1} | X_{1,\dots,n}, \theta^{i-1}) \\ &\propto m_{S_{t-1} \mapsto S_t}(s_t) \cdot T^{(i-1)}[s_{t-1}, s_t] \cdot P(S_{t-1} = s_{t-1} | X_{1,\dots,n}, \theta^{i-1}) \\ &\propto m_{S_{t-1} \mapsto S_t}(s_t) \cdot T^{(i-1)}[s_{t-1}, s_t] \\ &\quad m_{S_{t-2} \mapsto S_{t-1}}(s_{t-1}) \cdot m_{S_t \mapsto S_{t-1}}(s_{t-1}) \cdot E^{(i-1)}[s_{t-1}, X_{t-1}] \end{aligned}$$

BAUM WELCH ALGORITHM

Initialize T^0, E^0 probability tables

For $i = 1$ to convergence

- E-step:
 - Run Forward-Backward algorithm and compute messages
 - For every t compute $P(S_t = s_t, S_{t-1} = s_{t-1} | X_{1,\dots,n}, \theta^{i-1})$ and $P(S_t = s_t | X_{1,\dots,n}, \theta^{i-1})$ as in previous slides
- M-step:

$$\forall u, v \quad T^{(i)}[u, v] = \frac{\sum_{t=2}^n P(S_t = v, S_{t-1} = u | X_{1,\dots,n}, \theta^{i-1})}{\sum_{t=2}^n P(S_{t-1} = u | X_{1,\dots,n}, \theta^{i-1})}$$

$$\forall v, e \quad E^{(i)}[v, e] = \frac{\sum_{t=1}^n P(S_t = v | X_{1,\dots,n}, \theta^{i-1}) \cdot \mathbf{1}_{X_t=e}}{\sum_{t=1}^n P(S_t = v | X_{1,\dots,n}, \theta^{i-1})}$$

Inference for general BN

BAYESIAN NETWORKS

- Directed acyclic graph (DAG): $G = (V, E)$
- Joint distribution P_θ over X_1, \dots, X_n that factorizes over G :

$$P_\theta(X_1, \dots, X_n) = \prod_{i=1}^n P_\theta(X_i | \text{Parent}(X_i))$$

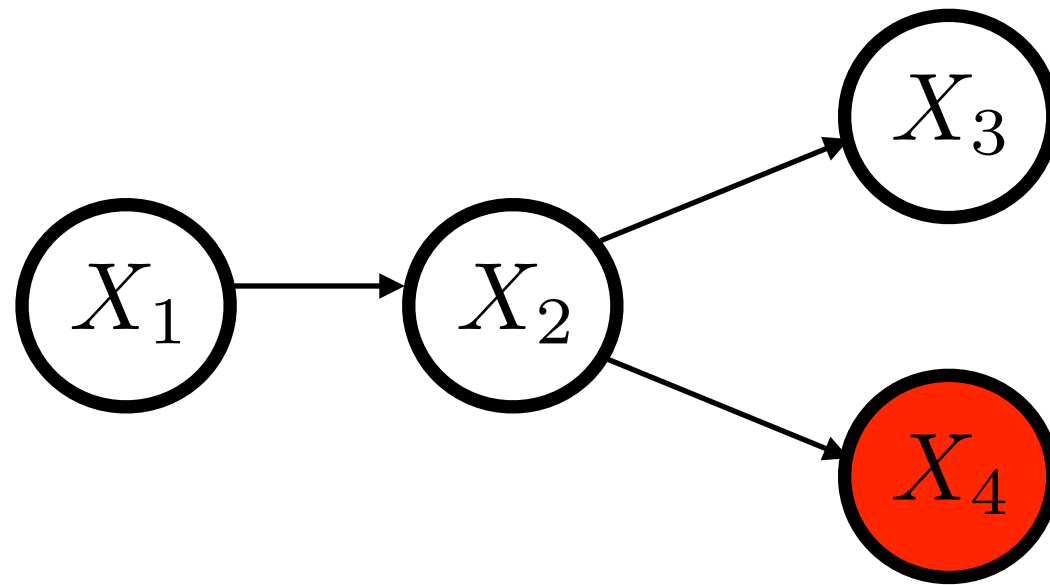
- Hence Bayesian Networks are specified by G along with CPD's over the variables (given their parents)

VARIABLE ELIMINATION: EXAMPLES

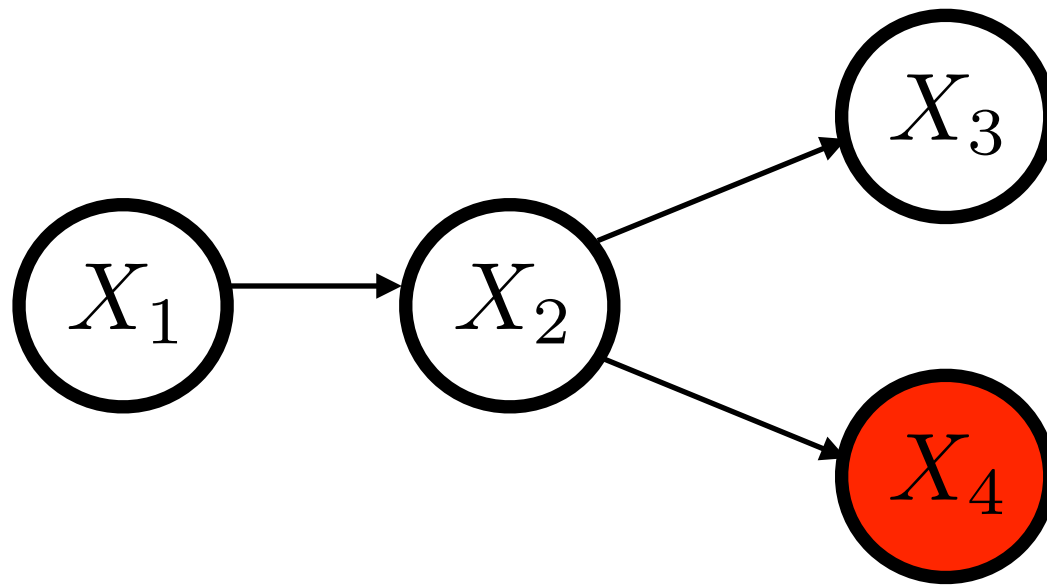
- Marginals are enough:

$$P(X_j = x_j, X_k = x_k | X_i = x_i, X_h = x_h) = \frac{P(X_j = x_j, X_k = x_k, X_i = x_i, X_h = x_h)}{P(X_i = x_i, X_h = x_h)}$$

VARIABLE ELIMINATION: EXAMPLES



VARIABLE ELIMINATION: EXAMPLES



$$\begin{aligned} P(X_4) &= \sum_{x_1} \sum_{x_2} \sum_{x_3} P(X_1 = x_1, X_2 = x_2, X_3 = x_3, X_4) \\ &= \sum_{x_1} \left(P(X_1 = x_1) \sum_{x_2} \left(P(X_2 = x_2 | X_1 = x_1) P(X_4 | X_2 = x_2) \left(\sum_{x_3} P(X_3 = x_3 | X_2 = x_2) \right) \right) \right) \\ &= \sum_{x_1} \left(P(X_1 = x_1) \left(\sum_{x_2} P(X_2 = x_2 | X_1 = x_1) P(X_4 | X_2 = x_2) \right) \right) \end{aligned}$$

VARIABLE ELIMINATION: BAYESIAN NETWORK

Initialize **List** with conditional probability distributions

Pick an order of elimination I for remaining variables

For each $X_i \in I$

 Find distributions in **List** containing variable X_i and remove them

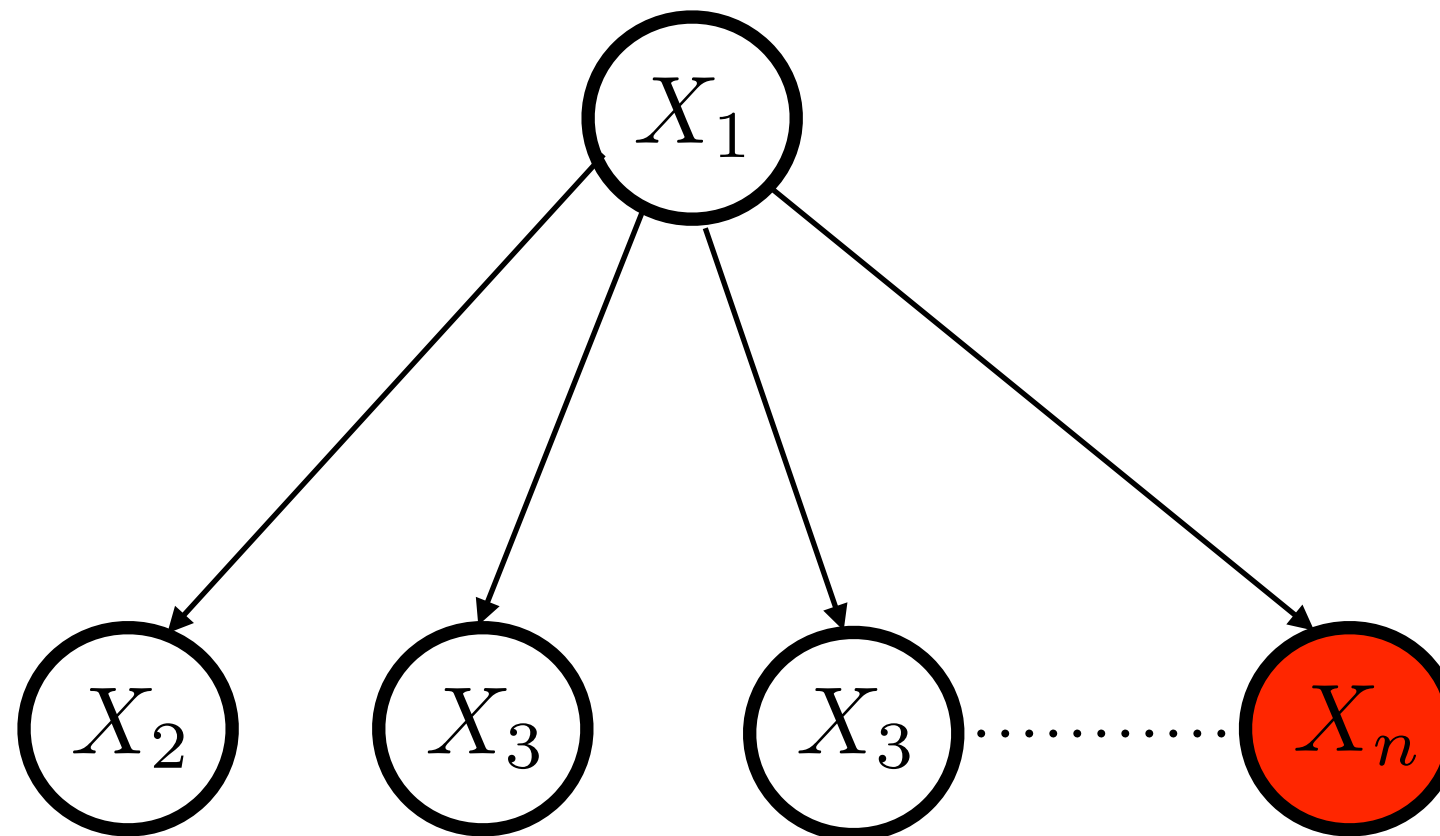
 Define new distribution as the sum (over values of X_i) of the product of these distributions

 Place the new distribution on **List**

End

Return **List**

VARIABLE ELIMINATION: ORDER MATTERS



Right order: $O(n)$

Wrong order: $O(2^n)$

MESSAGE PASSING

- Often we need more than one marginal computation
- Over variables we need marginals for, there are many common distributions/potentials in the list
- Can we exploit structure and compute these intermediate terms that can be reused?

Eg. forward backward algorithm for HMM